

A mixed hidden Markov model for analyzing somatic cell scores

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Outline

- Introduction
- Mixed Hidden Markov model
- Simulation
- Conclusions

1. Introduction

Biological marker

Accuracy of available marker

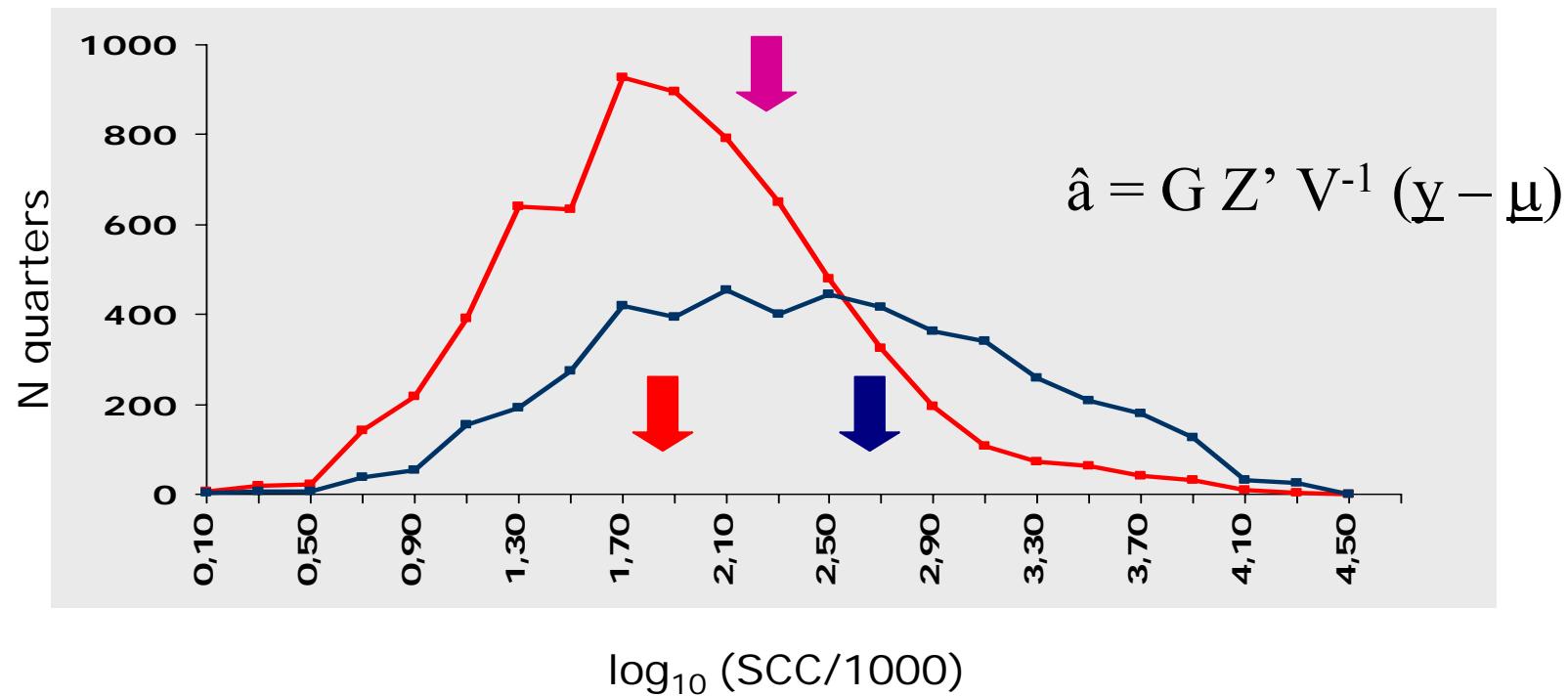
Resistance/tolerance

1a. Biological marker

- Reduction of bovine mastitis prevalence
(eg. breeding for disease resistance)
 - Early detection of mammary infection (IMI)
 - Biomarker = objective indicator of disease state
(eg., SCC, M-SAA3, Hp, LDH, NAGAse, ...)
 - Surrogate endpoint = substitute for disease endpoint
(eg. early predictor of infection, survival, or clinical signs)
- mathematical models to estimate pr(IMI)

1b. Accuracy

Frequency distribution of SCC for IMI- or IMI+ quarters



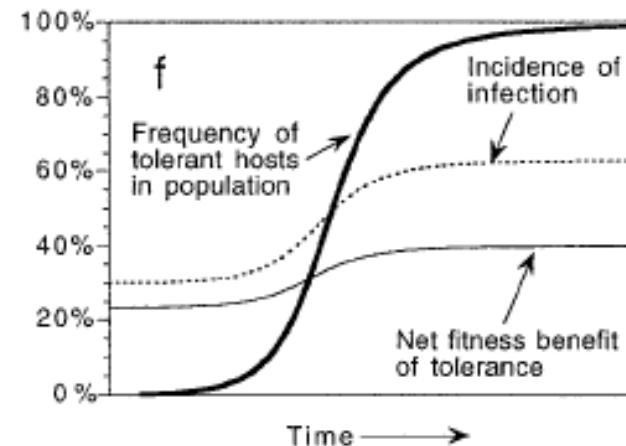
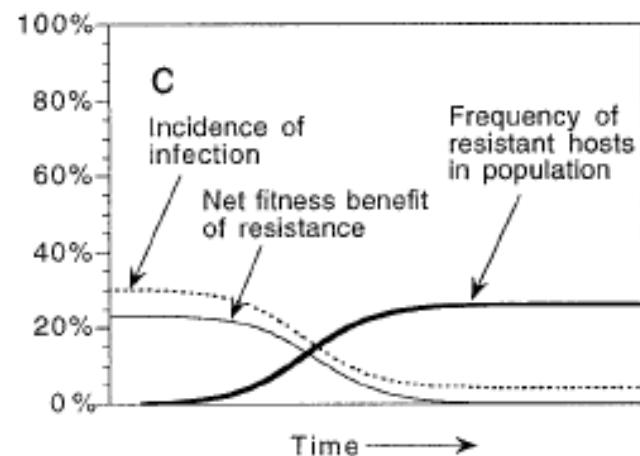
Imperfect detectability

→ misleading info on transmission dynamics
(prevalence, incidence, association with disease, ...)

1c. Resistance/ tolerance

B. A. ROY¹ AND J. W. KIRCHNER² *Evolution*, 54(1), 2000, pp. 51–63

- Resistance = low proba of infection
- Tolerance = little fitness loss after infection



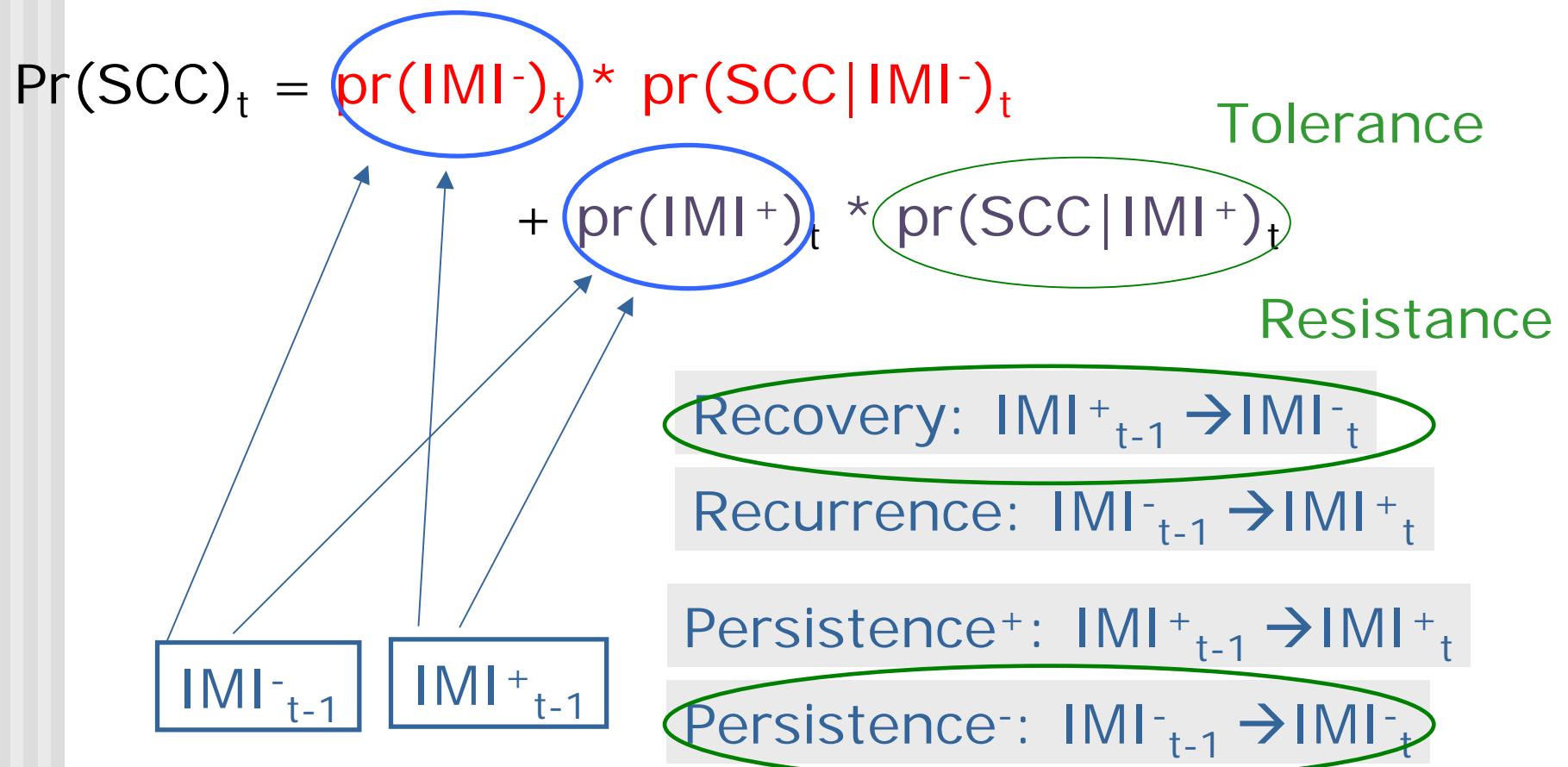
Breeding goals ?

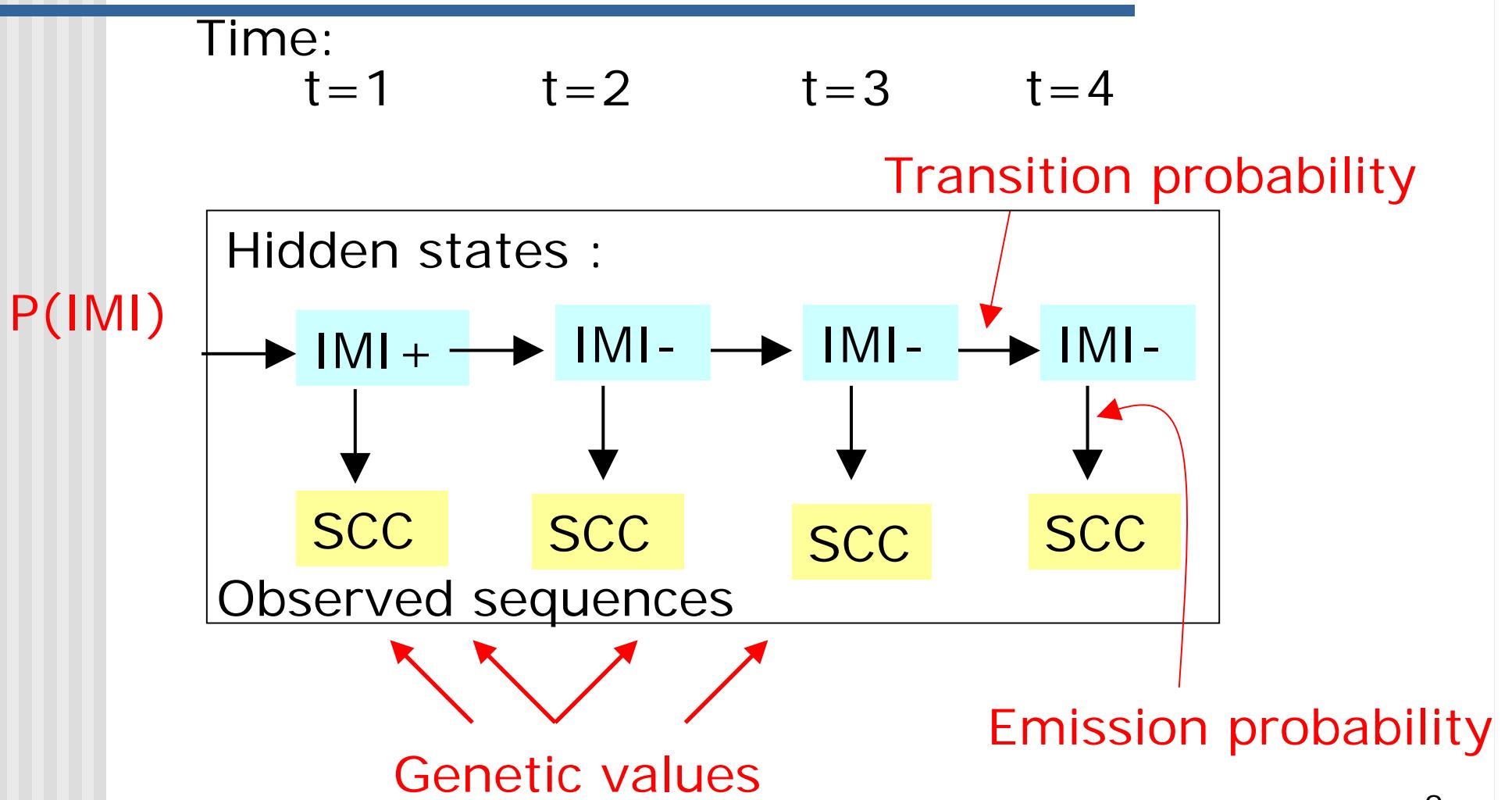
- + herd immunity
- natural selection
- disease spread
- + natural selection

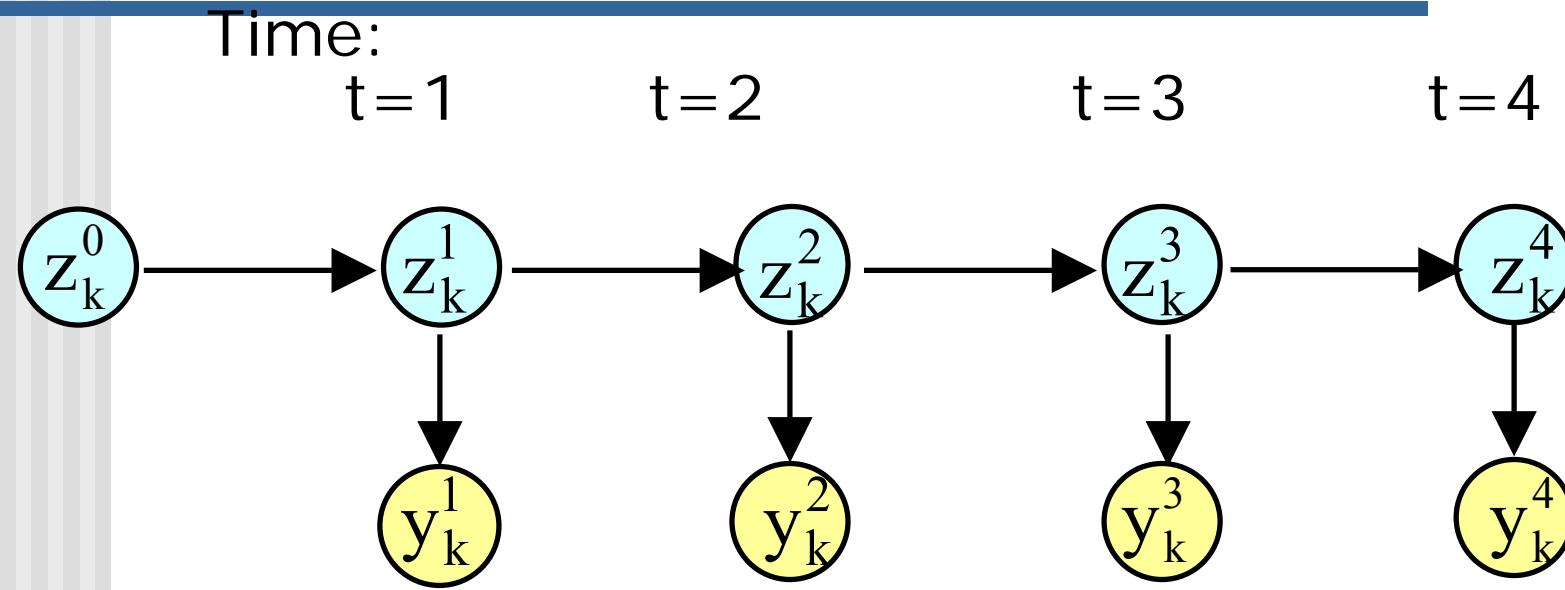
2. Mixed hidden Markov model

General formulation
Bayes estimation

2a. Formulation







y_k^t = value of biomarker at t^{th} time on k^{th} cow

$$z_k^t = 0 \text{ if } |M| -$$

$$Z_k^t = 1 \text{ if } |MI| +$$

- Output independence: observations are independent given the unknown IMI state

$$p(y_k^t | z_k^t, y_k^{t-1}, y_k^{t-2}, \dots) = p(y_k^t | z_k^t)$$

- Time is discrete
- Markov property: the next state depends only on the current state

$$p(z_k^{t+1} | z_k^t, z_k^{t-1}, \dots, z_k^1) = p(z_k^{t+1} | z_k^t)$$

“Conditioned on the present, the past & future are independent”

- Stationarity: transition probabilities are time invariant

$$p(z_k^{t+1} = i | z_k^t = j) = a_k^{ij}$$

$$p(\underline{y} | \underline{\mu}_0, \underline{\mu}_1, \sigma_0^2, \sigma_1^2, \underline{a}, \underline{z}) \sim N(M_0 \underline{\mu}_0 + M_1 \underline{\mu}_1 + Z \underline{a}, R)$$

\underline{y} = (NT X 1) vector of data

\underline{z} = (NT X 1) vector of hidden states

$\underline{\mu}_0$ = (T X 1) vector of fixed effects for data on a IMI- cow

$\underline{\mu}_1$ = (T X 1) vector of fixed effects for data on a IMI+ cow

\underline{a} = (N X 1) vector of random additive genetic effects

M_0 = (NT X T) matrix with elements = 1 if $z_k^t = 0$

M_1 = (NT X T) matrix with elements = 1 if $z_k^t = 1$

Z = (NT X N) incidence matrix relating \underline{a} to \underline{y} .

$R \sim J_0 \sigma_0^2 + J_1 \sigma_1^2$ J_i = (NT X NT) matrix with elements = 1 if $z_k^t = i$

2b. Estimation

Prior distributions

$$\underline{\mu}_0 \sim N(\underline{m}_0, I s_0^2)$$

$$\underline{\mu}_1 \sim N(\underline{m}_1, I s_1^2)$$

$$\underline{a} \sim N(0, A \sigma_a^2)$$

$$\sigma_a^2 \sim \chi^{-2}(s_a^2)$$

$$\sigma_0^2 \sim \chi^{-2}(s_0^2)$$

$$\sigma_1^2 \sim \chi^{-2}(s_1^2)$$

$$z_k^0 \sim Br(\lambda_k)$$

$$(z_k^t = 0 | z_k^{t-1} = 0) \sim Br(\pi_k^{00})$$

$$(z_k^t = 1 | z_k^{t-1} = 0) \sim Br(\pi_k^{01})$$

$$\lambda_k, \pi_k^{00}, \pi_k^{01} \sim Un[0,1]$$

Joint posterior distribution

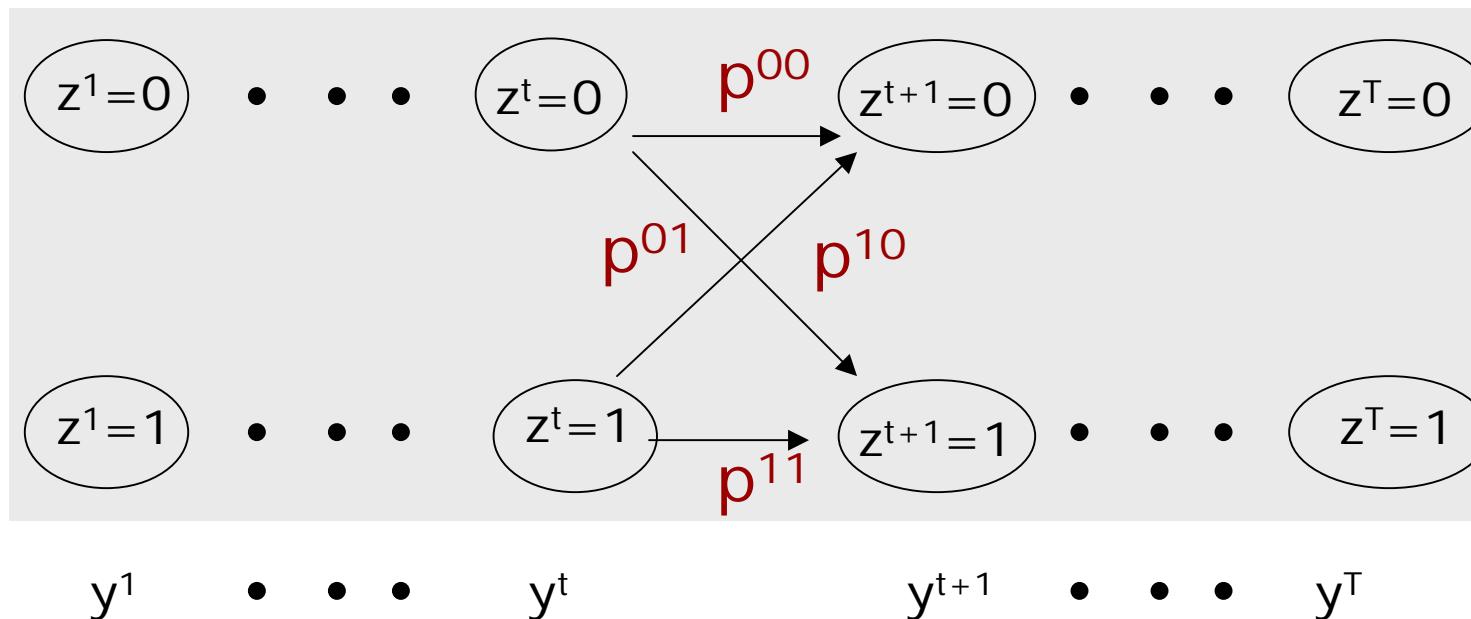
$$\theta = (\mu_0, \mu_1, \sigma_a^2, \sigma_0^2, \sigma_1^2, \underline{\pi}^{00}, \underline{\pi}^{01}, \underline{z}, \underline{a}, \underline{\lambda})$$

$$p(\theta | \underline{y}) = \sum_{\underline{z}} p(\underline{y} | \theta) p(\theta)$$

nb hidden sequences per cow = 2^T
→ total number of operations $\sim 4N^T$

Use of the trellis structure of HMM
→ Forward-backward algorithm

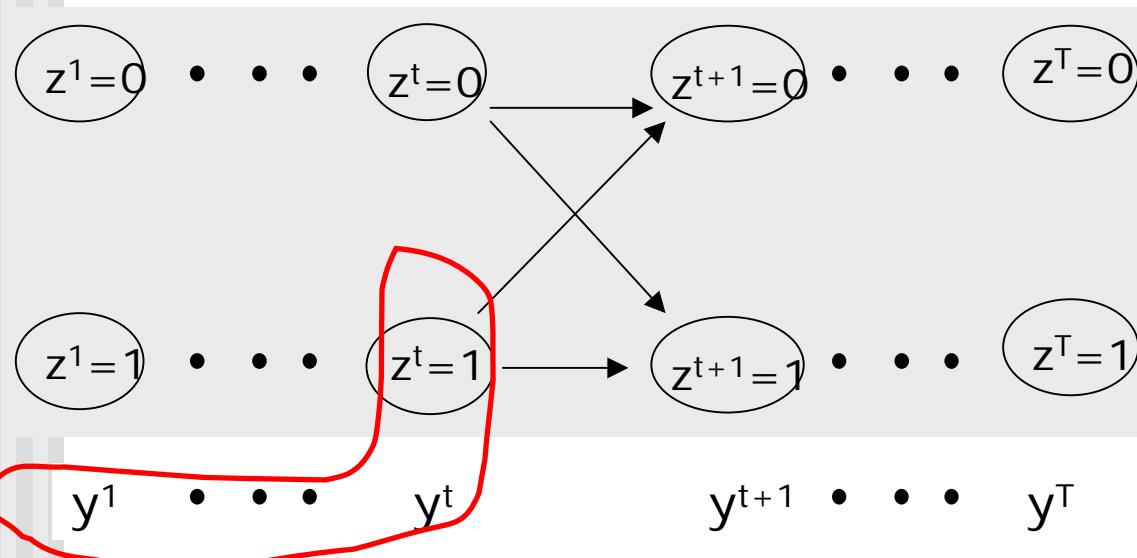
Forward-backward algorithm



The algorithm takes on the order of $4T$ computations

- **Forward probabilities** : proba that, given H, at time t, the state is i and the sequence of partial observation ($y_1 \dots y_t$) has been generated

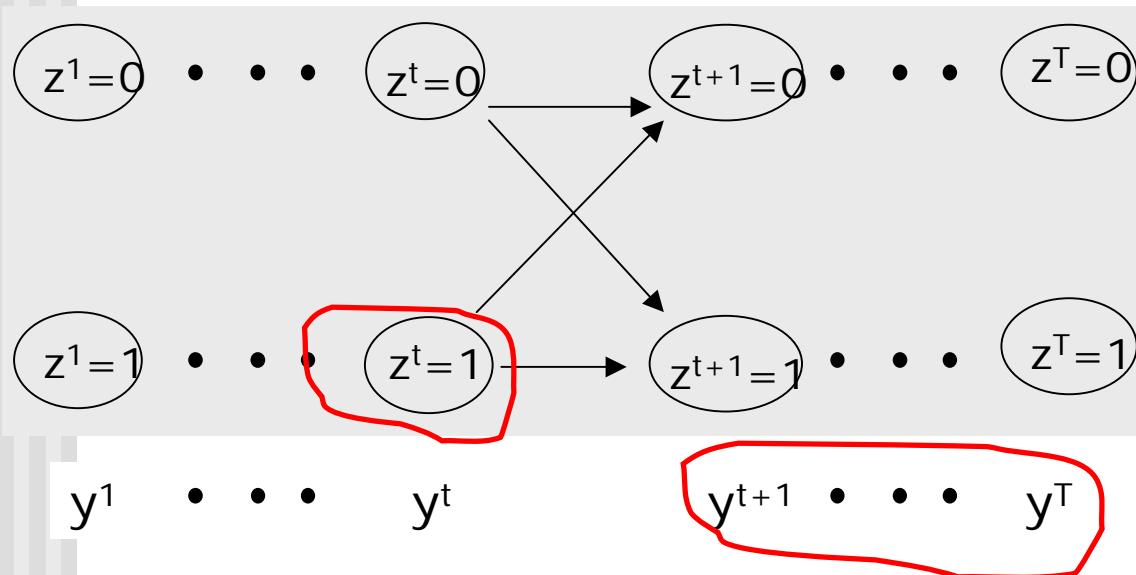
$$\alpha_k^t(i) = p(y_k^1 \dots y_k^t, z_k^t = i)$$



3 steps
 induction ($t=0$)
 recursion (increasing t)
 termination ($t=T$)

- **Backward probabilities** : proba that, given H and given the state i at time t, a sequence of partial observation ($y_{t+1} \dots y_T$) has been generated

$$\beta_k^t(i) = p(y_k^T \dots y_k^{t+1} | z_k^t = i)$$



3 steps
 induction ($t=T$)
 recursion (decreasing t)
 termination ($t=1$)

Fully conditional distributions

$$\mu_i^t \sim N\left(\frac{s_i^2 \sum_k^N (y_k^t - a_k) I_k^{i,t} + m_i \sigma_i^2}{(s_i^2 \sum_k^N n_{i,k}^t) + \sigma_i^2}, \frac{s_i^2 \sigma_i^2}{(s_i^2 \sum_k^N n_{i,k}^t) + \sigma_i^2}\right)$$

$$\underline{a} \sim N(\tilde{\theta}_1, C_{11}^{-1}) \quad \tilde{\theta}_1 = C_{11}^{-1} [r_1 - C_{12} \theta_2]$$

$$(\sigma_a^2) \sim \chi_{N+v}^{-2} \quad (\underline{a}' A^{-1} \underline{a} + v s_a^2)$$

$$(\sigma_i^2) \sim \chi_{N_i+v}^{-2} \quad [v s_i^2 + (\underline{y} - M_i \underline{\mu}_i - Z \underline{a})' J_i (\underline{y} - M_i \underline{\mu}_i - Z \underline{a})]$$

$$\lambda_k \sim \text{beta}(I_k^{0,1} + 1; I_k^{1,1} + 1)$$

$$\pi_k^{00} \sim \text{beta}(n_k^{00} + 1, n_k^{10} + 1) \quad \pi_k^{01} \sim \text{beta}(n_k^{01} + 1, n_k^{11} + 1)$$

$$\underline{z}_k^1 \sim \text{Br}(\zeta_{0,k}^1) \quad \zeta_{0,k}^0 = p[z_k^0 = 0 \cap \underline{y}_k] = \alpha_k^1(0) \beta_k^1(0)$$

It is the probability of being healthy at start with the observation sequence $y_1 y_2 \dots y_T$.

$$\begin{aligned} \underline{z}_k^t &\sim \text{Br}(\xi_{ij,k}^t) & \xi_{ij,k}^t &= p[z_k^t = i | z_k^{t-1} = j, \underline{y}_k] \\ && &= \frac{\alpha_k^{t-1}(j) \pi_k^{ij} \beta_k^t(i) \text{pr}(y_k^t | z_k^t = i)}{\alpha_k^{t-1}(j) \beta_k^{t-1}(j)} \end{aligned}$$

It is the probability of being in state i at time t given state j at time $t-1$ and the observation sequence $y_1 y_2 \dots y_T$.

4. Simulation

Survey of SCC

Pathogen

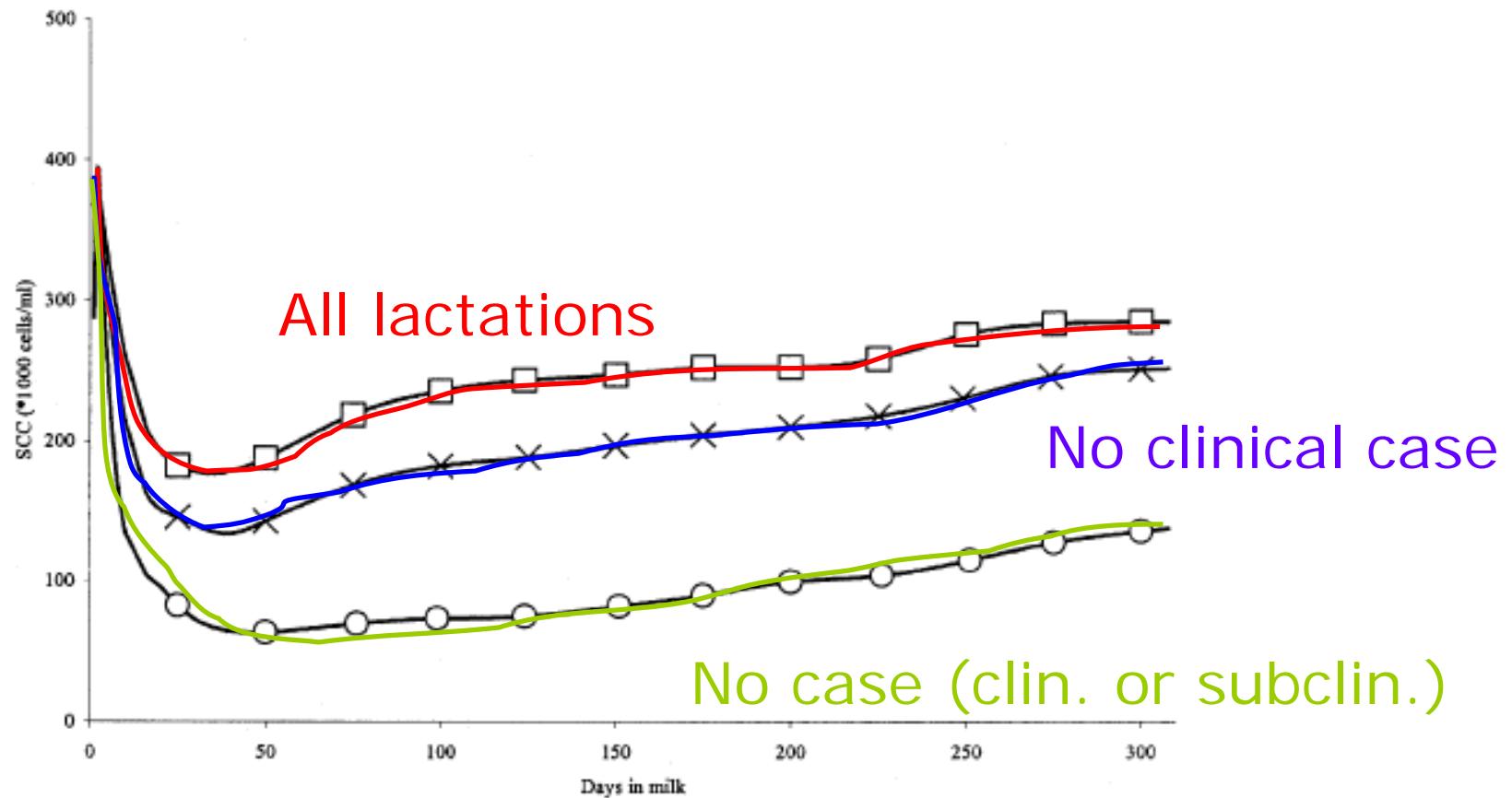
Severity of response

Genetics

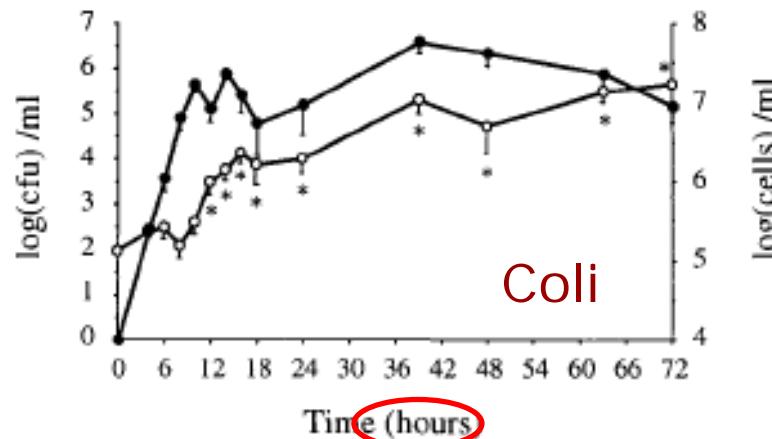
Data sets

Accuracy

4a. Survey (de Haas et al., 2002, 2004)

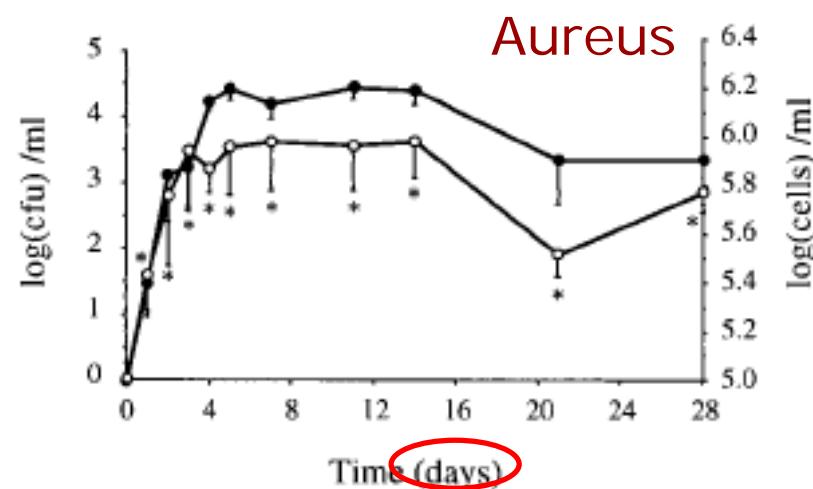


4b. Pathogen



Coli

Time (hours)



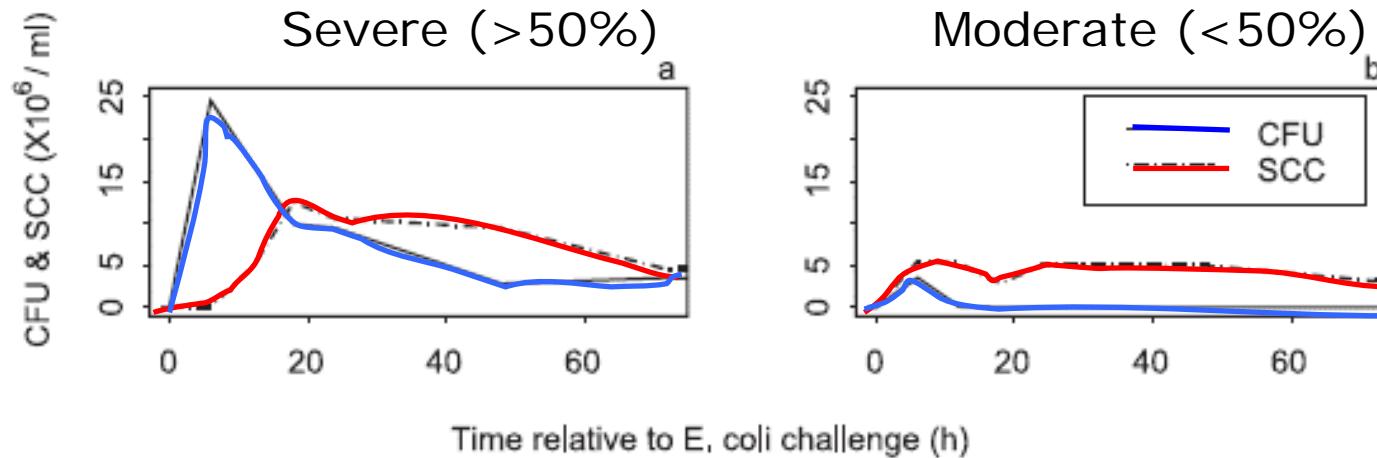
Aureus

Time (days)

CÉLINE RIOLLET, PASCAL RAINARD,* AND BERNARD POUTREL

CLINICAL AND DIAGNOSTIC LABORATORY IMMUNOLOGY, Mar. 2000, p. 161–167

4c. Severity of response



Jalil MEHRZAD^{a,b,c}, Luc DUCHATEAU^a, Christian BURVENICH^{a*}
Vet. Res. 36 (2005) 101–116

4d. Genetics

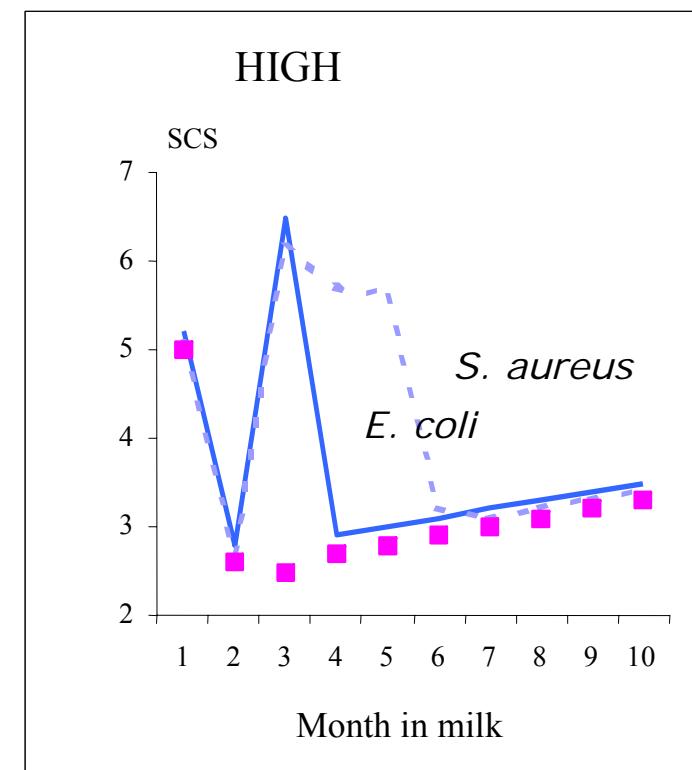
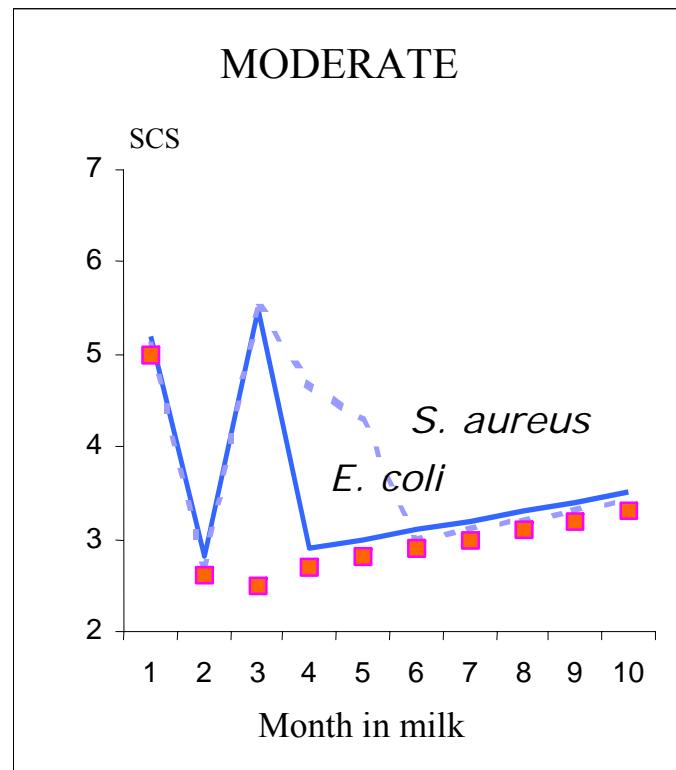
- 3 discrete generations with 400 cows per generation
- sires selected from 30 different bulls
- each cow was replaced by a daughter
- mating at random

- breeding values for base animals $\sim N(0, I \sigma_a^2)$
- additive variance of 0.15 or 0.25.
- breeding values for non-base animals $\sim N(\text{mid-parent}, \sigma_a^2/2)$
- No selection , no inbreeding

4e. Simulated data sets

μ_0^t for $t = 1$ to T

μ_1^t for $t = 1$ to T



Simulated data sets:

% infected cows = 20, 50%

% *E. coli* among infected cows = 0,50,100%

high and moderate responders: μ_1^t

$\sigma_0^2 = 1.0 \text{ or } 1.4 \rightarrow \text{residuals IMI-} \sim N(0, \sigma_0^2)$

$\sigma_a^2 = 0.15 \text{ or } 0.25$

Gibbs sampler

1000 iterations

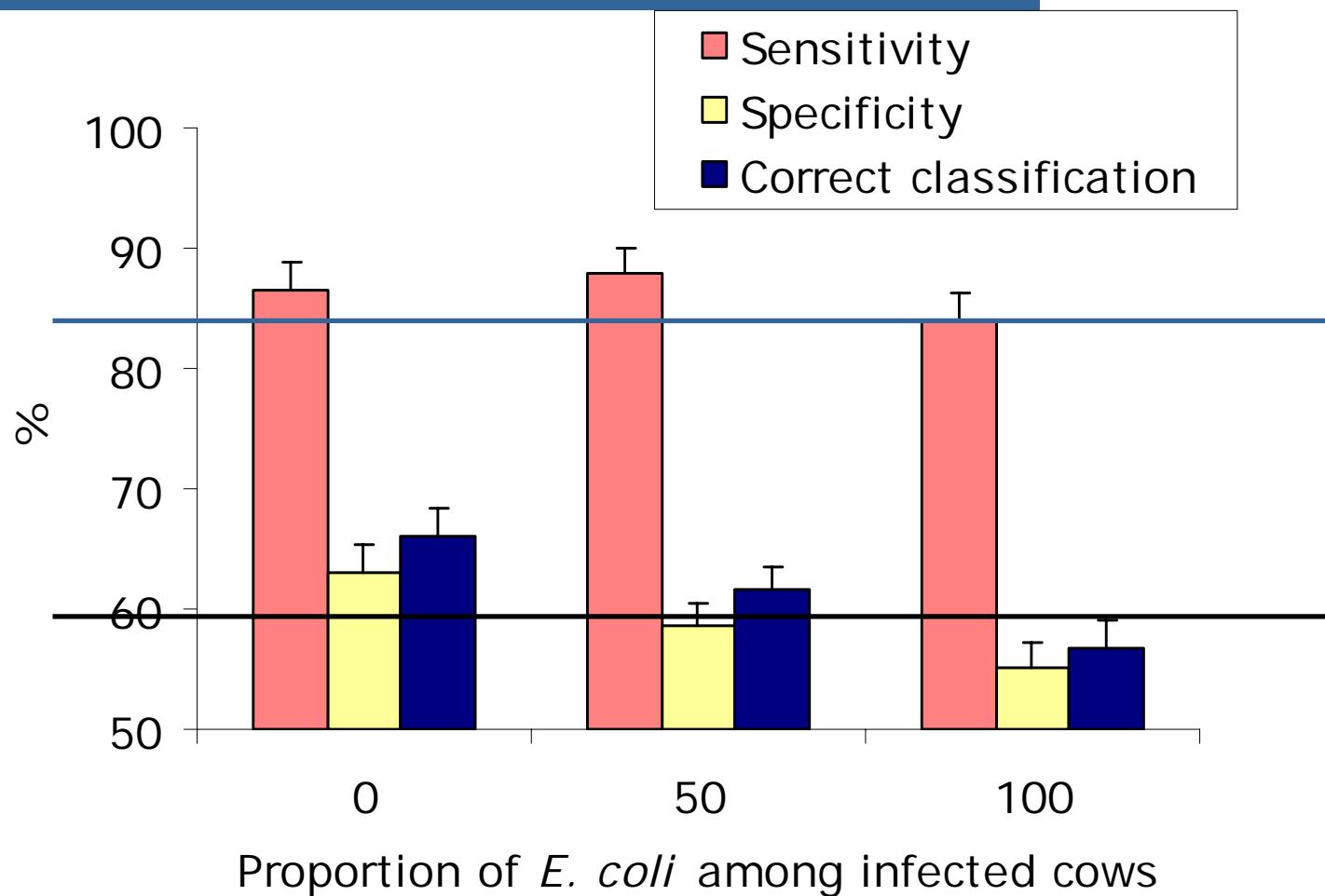
200 burn-in

10 replications

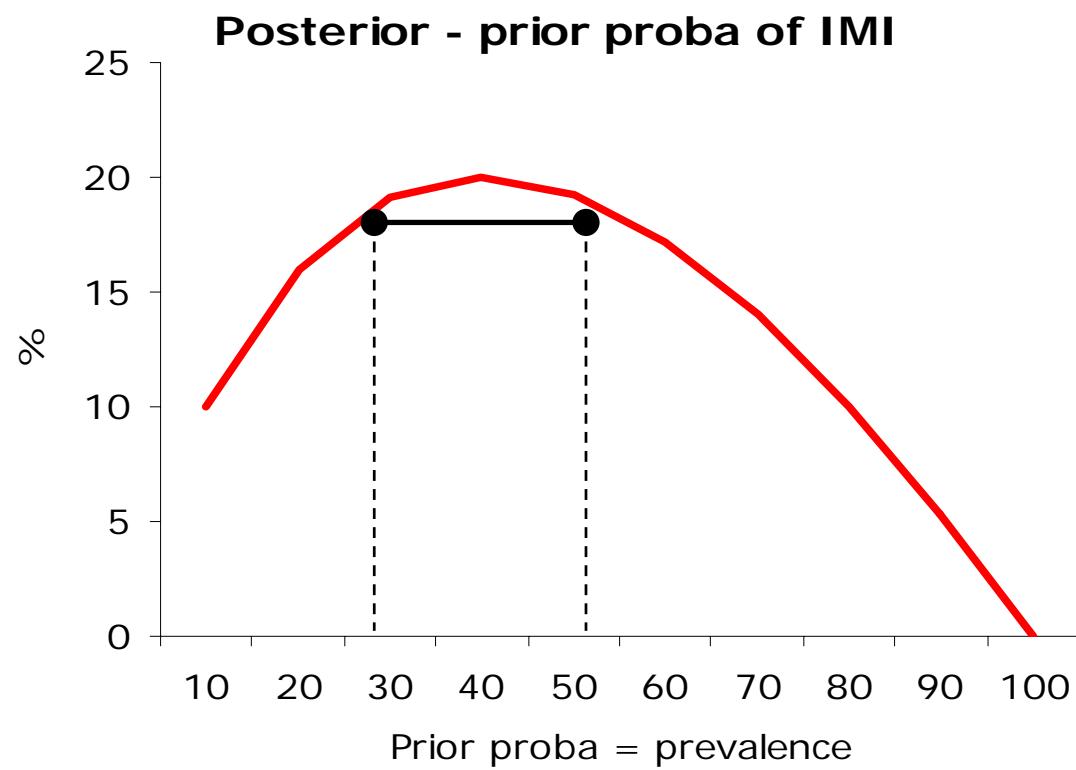
4f. Accuracy

- $\text{corr}(a, \hat{a})$
- differences : $\theta - \hat{\theta}$ for $\mu_0, \mu_1, \sigma_a^2, \sigma_0^2, \sigma_1^2$

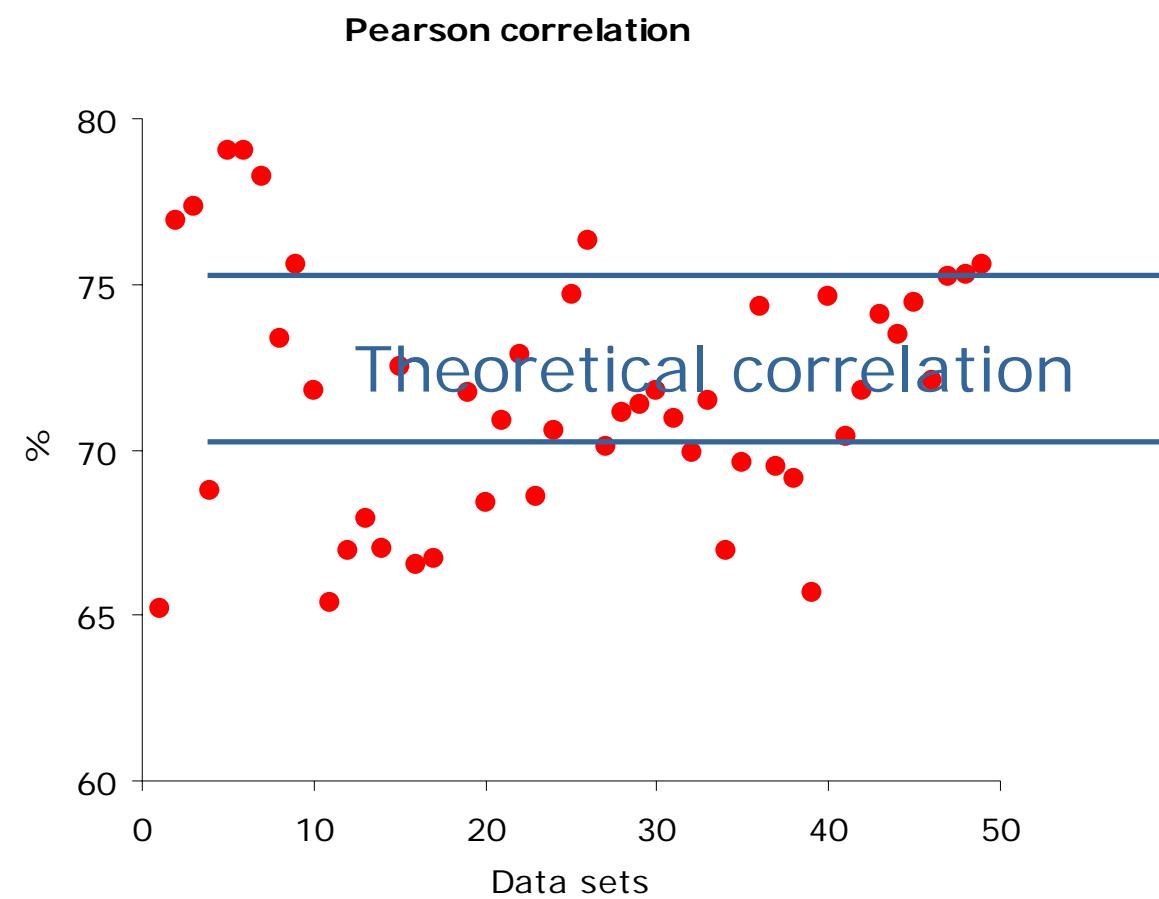
- sensitivity: $SE = \sum_{k=1,N} \sum_{t=1,T} p(\hat{z}_k^t = 1 | z_k^t = 1)$
- specificity: $SP = \sum_{k=1,N} \sum_{t=1,T} p(\hat{z}_k^t = 0 | z_k^t = 0)$
- proba of correct classification:
$$PCC = \sum_{k=1,N} \sum_{t=1,T} p[(z_k^t = 1 \cap \hat{z}_k^t = 1) \cup (z_k^t = 0 \cap \hat{z}_k^t = 0)]$$



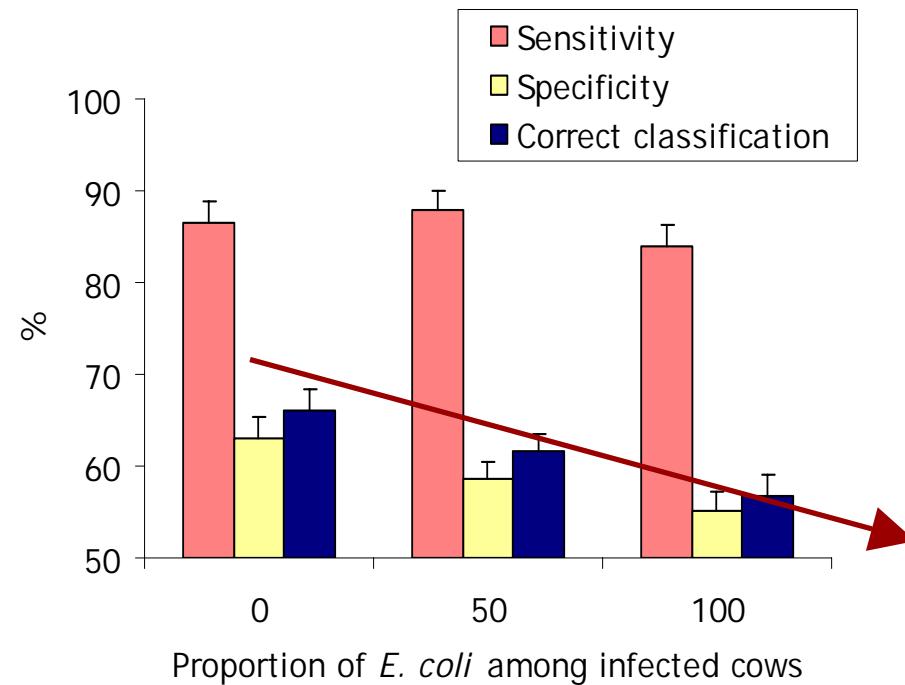
High SE = SNOUT

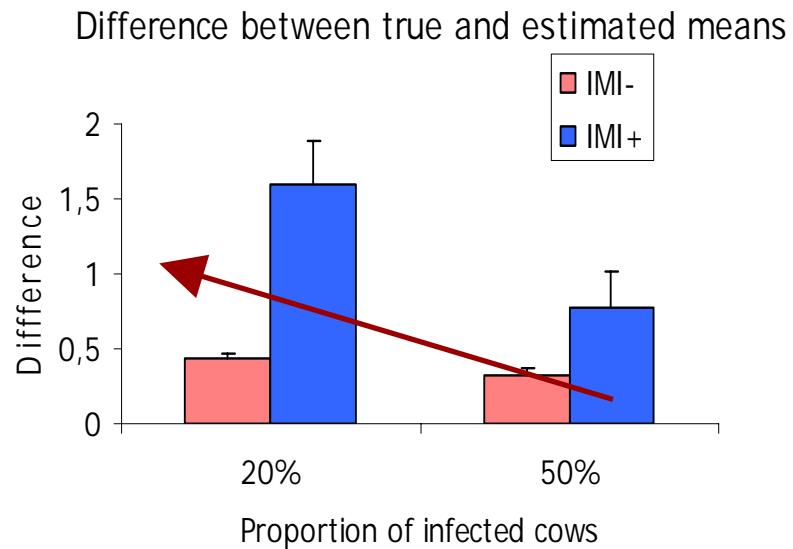


→ No further test



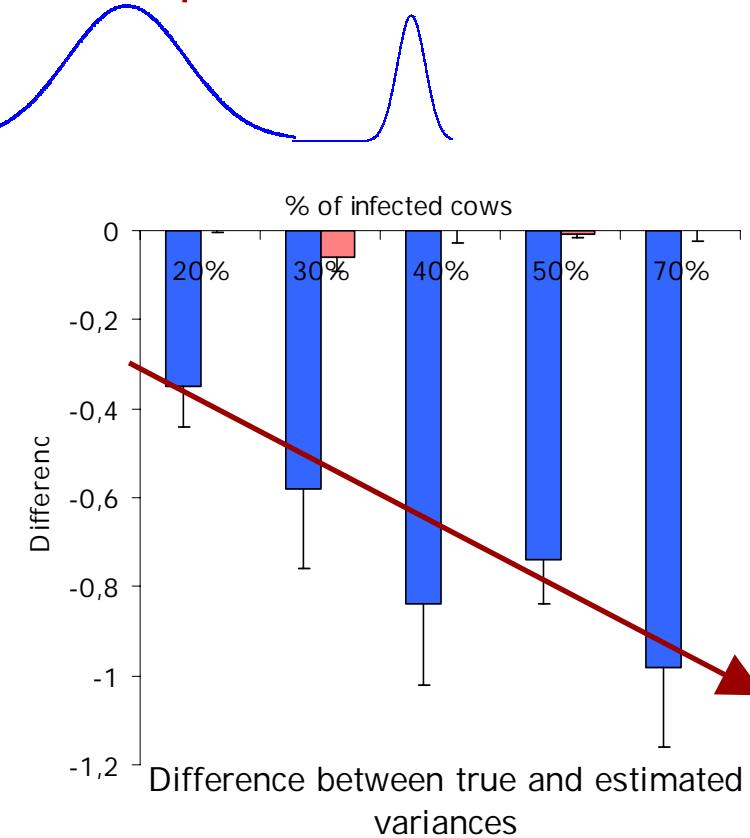
SP and PCC decreased when
% *E. coli* among infected increased





Biases increased when disease prevalence is high

Biases increased when disease prevalence is low



Conclusions

- Same amount of data
 - Increased accuracy of MLE
 - Resistance and tolerance
 - Transition probabilities

- Simplification of reality
 - Age, season, herd, ..
 - 'Isolated' proba of IMI-
 - Genetic relationship between cows for IMI



Genetic relationship among IMI and SCC

