

Modelling individual patterns of somatic cell scores to derive mammary infection status

J. Detilleux
University of Liège
Faculty of veterinary medicine

jdetilleux@ulg.ac.be



Outline

- Introduction
- Finite mixture model
- Hidden Markov model
- Simulation
- Conclusions

1. Introduction

Biological marker

Accuracy of available marker

Resistance/tolerance

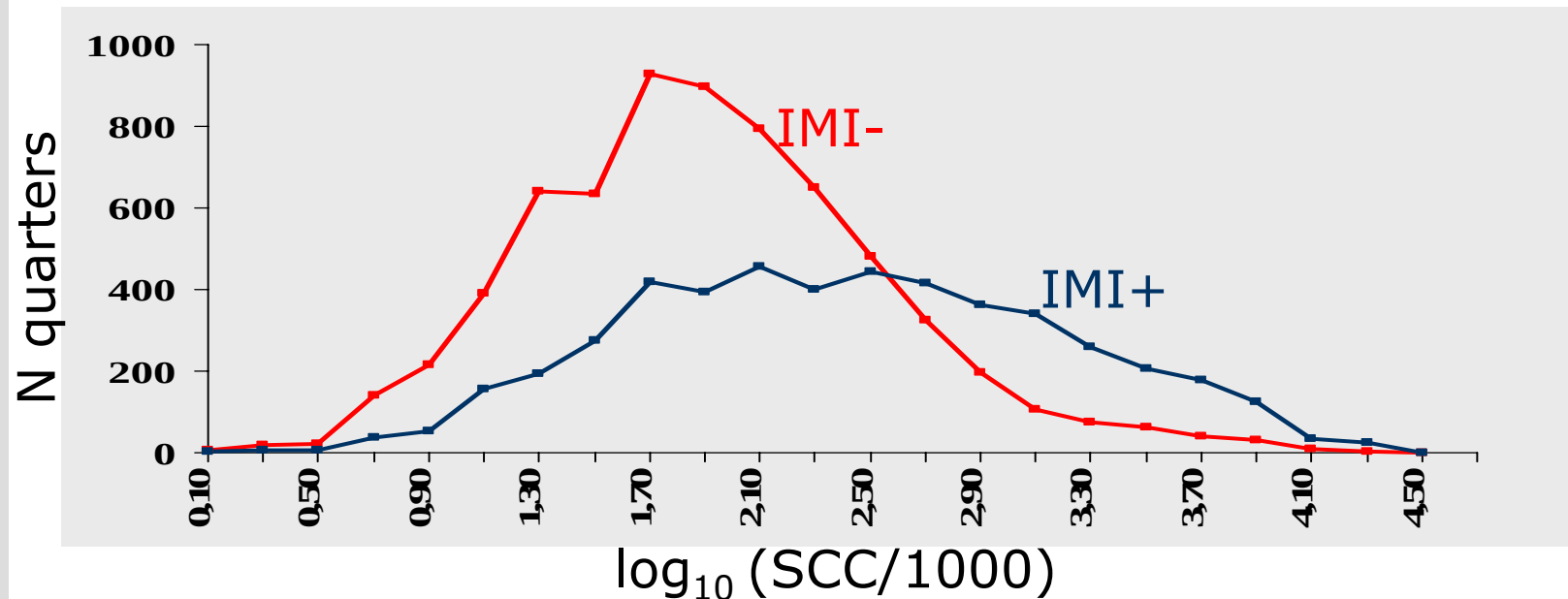
1a. Biological marker

- Reduction of bovine mastitis prevalence
Opposition to disease transmission
- Early detection of mammary infection (IMI)
 - Biomarker = objective indicator of disease state
(eg., SCC, M-SAA3, Hp, LDH, NAGase, ...)
 - Surrogate endpoint = substitute for disease endpoint
(eg. early predictor of infection, survival, or clinical signs)

→ mathematical models to estimate $\text{pr}(\text{IMI})$

1b. Accuracy

Frequency distribution of SCC for IMI- or IMI+ quarters



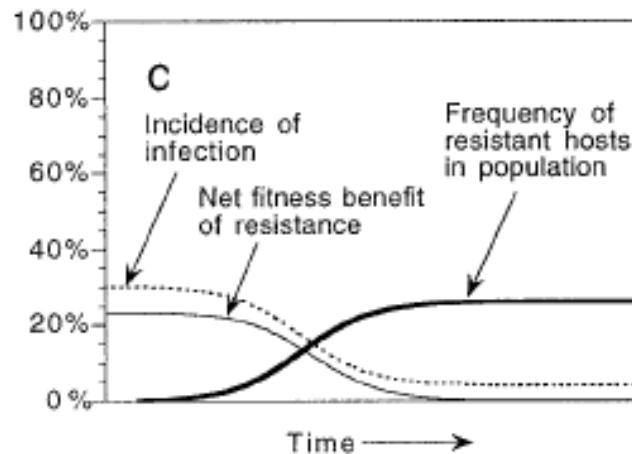
Imperfect detectability

→ misleading info on transmission dynamics
(prevalence, incidence, association with disease, ...)

1c. Resistance/ tolerance

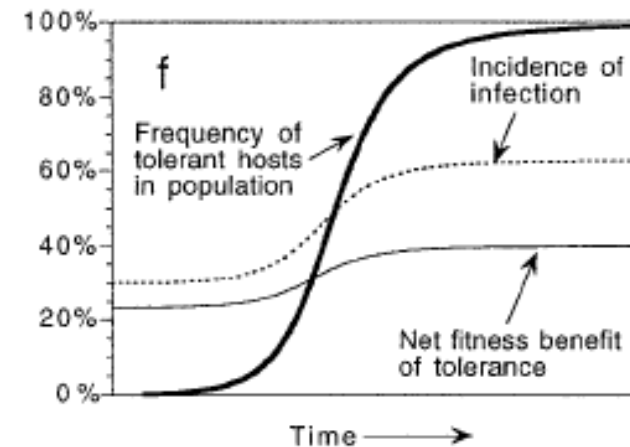
B. A. ROY¹ AND J. W. KIRCHNER² *Evolution*, 54(1), 2000, pp. 51-63

- Resistance = low proba of infection



- + herd immunity
- natural selection

- Tolerance = little fitness loss after infection



- disease spread
- + natural selection

2. Finite mixture model

General formulation

Likelihood

EM algorithm

2a. Formulation of FMM

$$\Pr(\text{SCC}) = \text{pr}(\text{SCC} \cap \text{IMI}^-) + \text{pr}(\text{SCC} \cap \text{IMI}^+)$$

$$\Pr(\text{SCC}) = \text{pr}(\text{IMI}^-) * \text{pr}(\text{SCC}|\text{IMI}^-) + [1 - \text{pr}(\text{IMI}^-)] * \text{pr}(\text{SCC}|\text{IMI}^+)$$

- Resistance
- Surrogate

- Tolerance
- Accuracy

FMM

Time:

t=1

t=2

t=3

t=4

Hidden states

IMI+

IMI-

IMI-

IMI-



SCC

SCC

SCC

SCC

Observed sequences

$P(\text{IMI})$

Emission probability:
 $P(\text{SCC}|\text{IMI})$

FMM

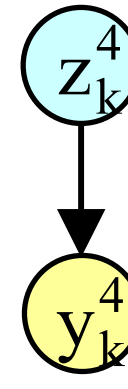
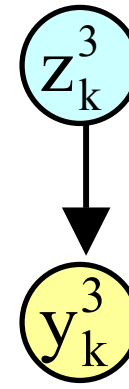
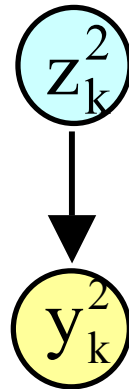
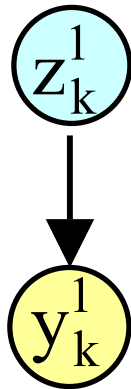
Time:

t=1

t=2

t=3

t=4



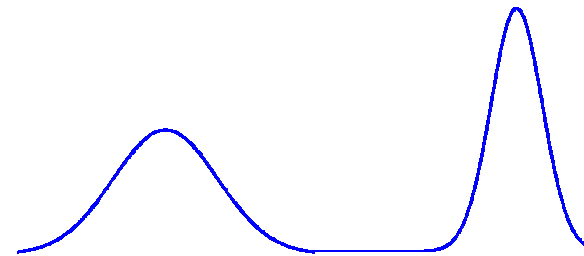
$$\begin{aligned} p(\underline{y}_k) &= p(y_k^1, y_k^2, \dots, y_k^T) \\ &= \prod_{t=1}^T p(y_k^t \mid z_k^t = 0)p(z_k^t = 0) + p(y_k^t \mid z_k^t = 1)p(z_k^t = 1) \end{aligned}$$

FMM

$$F = (\pi_k, \mu_0, \mu_1, \sigma_0^2, \sigma_1^2)$$

$$p(y_k^t | z_k^t = 0) \sim N(\mu_0^t, \sigma_0^2)$$

$$p(y_k^t | z_k^t = 1) \sim N(\mu_1^t, \sigma_1^2)$$

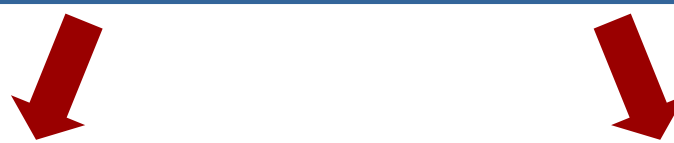


$$p(z_k^t = 0) \sim \text{Bi}(\pi_k)$$

cows are independent: $p(\underline{y}_1, \underline{y}_2, \dots, \underline{y}_N) = \prod_{k=1}^N p(\underline{y}_k)$

2b. Maximum likelihood estimation

Likelihood for one sequence (one cow):

$$p(\underline{y}_k | F) = p(y_k^1, y_k^2, \dots, y_k^T | F) = \sum_{\underline{z}} p(y_k^1, y_k^2, \dots, y_k^T | \underline{z}, F) \times p(\underline{z} | F)$$


$$\begin{aligned} & p(y_k^1, y_k^2, \dots, y_k^T | \underline{z}, F) \\ &= p(y_k^1, y_k^2, \dots, y_k^T | z_k^1, z_k^2, \dots, z_k^T, F) \\ &= p(y_k^1 | z_k^1, z_k^2, \dots, z_k^T, F) p(y_k^2 | y_k^1, z_k^1, z_k^2, \dots, z_k^T, F) \dots \\ &= p(y_k^1 | z_k^1, F) p(y_k^2 | z_k^2, F) \dots p(y_k^T | z_k^T, F) \\ &= \prod_{t=1}^T p(y_k^t | z_k^t, F) \end{aligned}$$

FMM

$$\begin{aligned} p(\underline{z} | F) &= p(z_k^1, z_k^2, \dots, z_k^T | F) \\ &= p(z_k^1 | F) p(z_k^2 | z_k^1, F) p(z_k^3 | z_k^2, z_k^1, F) \dots p(z_k^T | z_k^{T-1}, z_k^{T-2}, \dots, z_k^1, F) \\ &= p(z_k^1 | F) p(z_k^2 | F) p(z_k^3 | F) \dots p(z_k^T | F) = \prod_{t=1}^T p(z_k^t | F) \end{aligned}$$


$$\begin{aligned} p(y_k^1, y_k^2, \dots, y_k^T | F) &= \sum_{\underline{z}} \prod_{t=1}^T p(y_k^t | z_k^t, F) p(z_k^t | F) \\ &= \prod_{t=1}^T p(y_k^t | z_k^t = 0, F) p(z_k^t = 0 | F) + \prod_{t=1}^T p(y_k^t | z_k^t = 1, F) p(z_k^t = 1 | F) \end{aligned}$$



Likelihood for N cows:

$$p(\underline{y} | F) = \prod_{k=1}^N p(\underline{y}_1, \dots, \underline{y}_N | F)$$

2c. EM algorithm

E step:

$$Q^{(p-1)} = E[\log[p(\underline{y}_1, \dots, \underline{y}_N; \underline{z}_1, \dots, \underline{z}_{N_k}^T | \Theta) | \underline{y}_1, \dots, \underline{y}_N; \Theta^{(p-1)}]]$$

$$= \sum_{k=1}^N \sum_{t=1}^T \left[E[z_k^t = 0 | \underline{y}_k] \{ \log(\pi) + \log p(y_k^t | z_k^t = 0) \} \right. \\ \left. + E[z_k^t = 1 | \underline{y}_k] \{ \log(1 - \pi) + \log p(y_k^t | z_k^t = 1) \} \right]$$


$$E[z_k^t = i | \underline{y}_k] = \frac{\pi_k N(\mu_i^t, \sigma_i^2)}{\pi_k N(\mu_0^t, \sigma_0^2) + (1 - \pi_k) N(\mu_1^t, \sigma_1^2)} = \zeta_{i,k}^t$$

It is the probability of being in state i at time t , given the SCS sequence $y_1 y_2 \dots y_T$.

FMM

M step: $\Theta^{(p)} = \operatorname{argmax} Q^{(p-1)}$

$$\hat{\pi}_k = \frac{\sum_{t=1, T} \zeta_{0, k}^t}{T}$$

= expected frequency in state IMI- for the cow

$$\mu_i^t = \frac{\sum_{k=1, N} \zeta_{i, k}^t y_k^t}{\sum_{k=1, N} \zeta_{i, k}^t}$$

$$\hat{\sigma}_i^2 = \frac{\sum_{k=1, N} \sum_{t=1, T} \zeta_{i, k}^t (y_k^t - \hat{\mu}_i^t)^2}{\sum_{k=1, N} \sum_{t=1, T} \zeta_{i, k}^t}$$

= weighted mean and variance

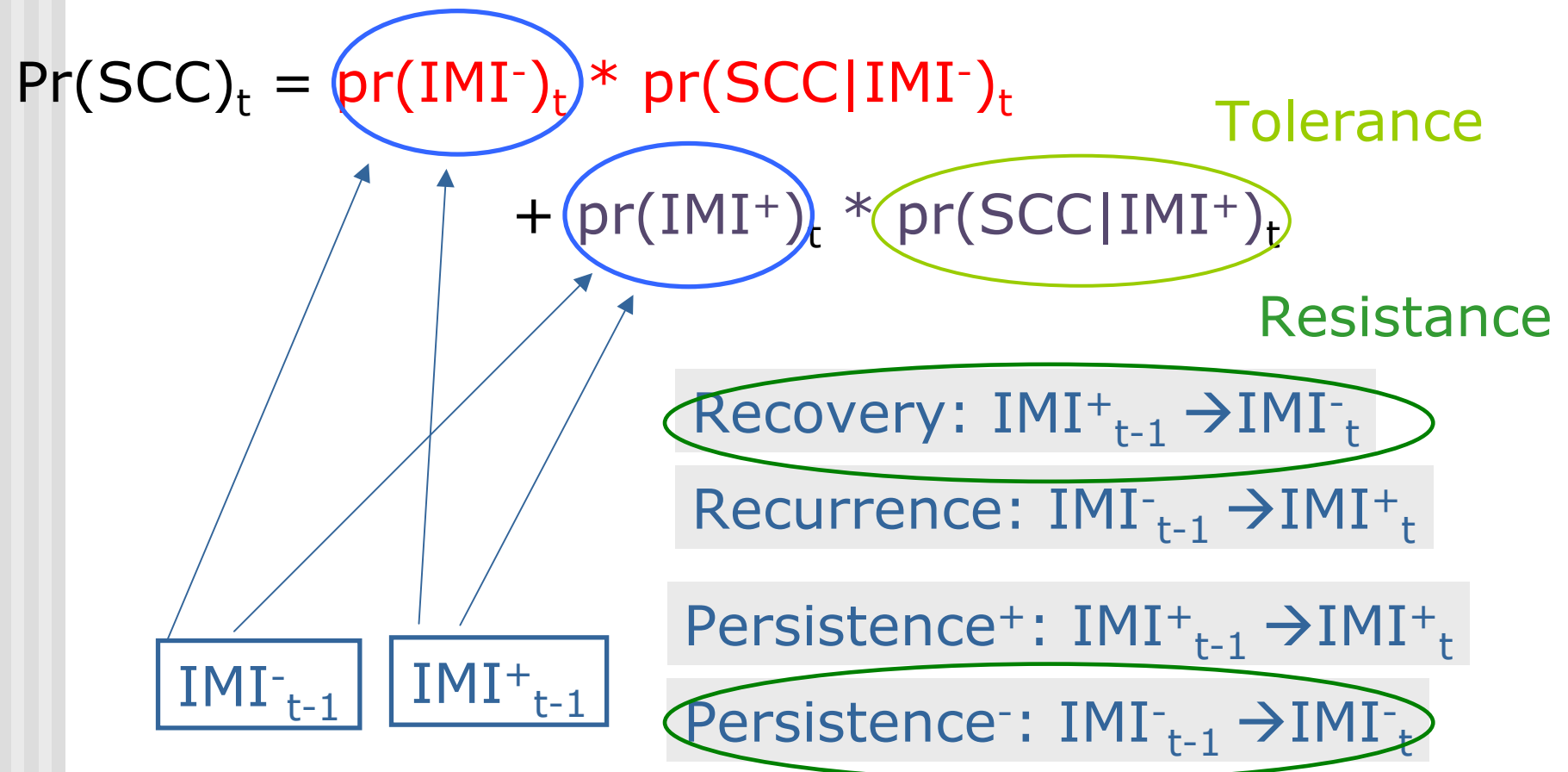
3. Hidden Markov model

General formulation

Likelihood

EM algorithm

2a. Formulation



Time:

t=1

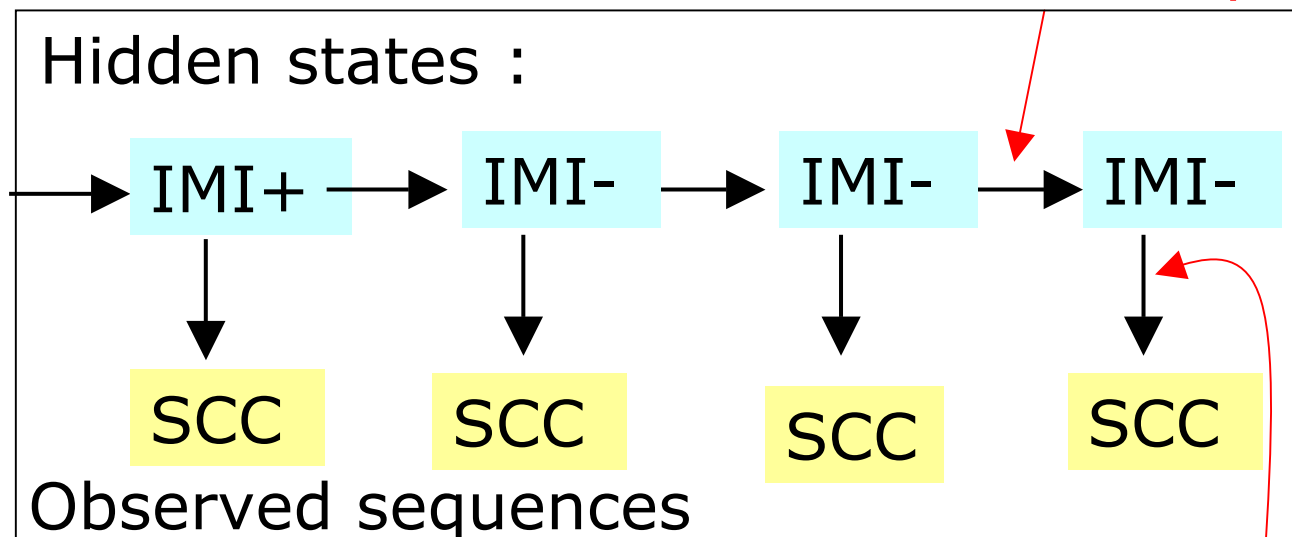
t=2

t=3

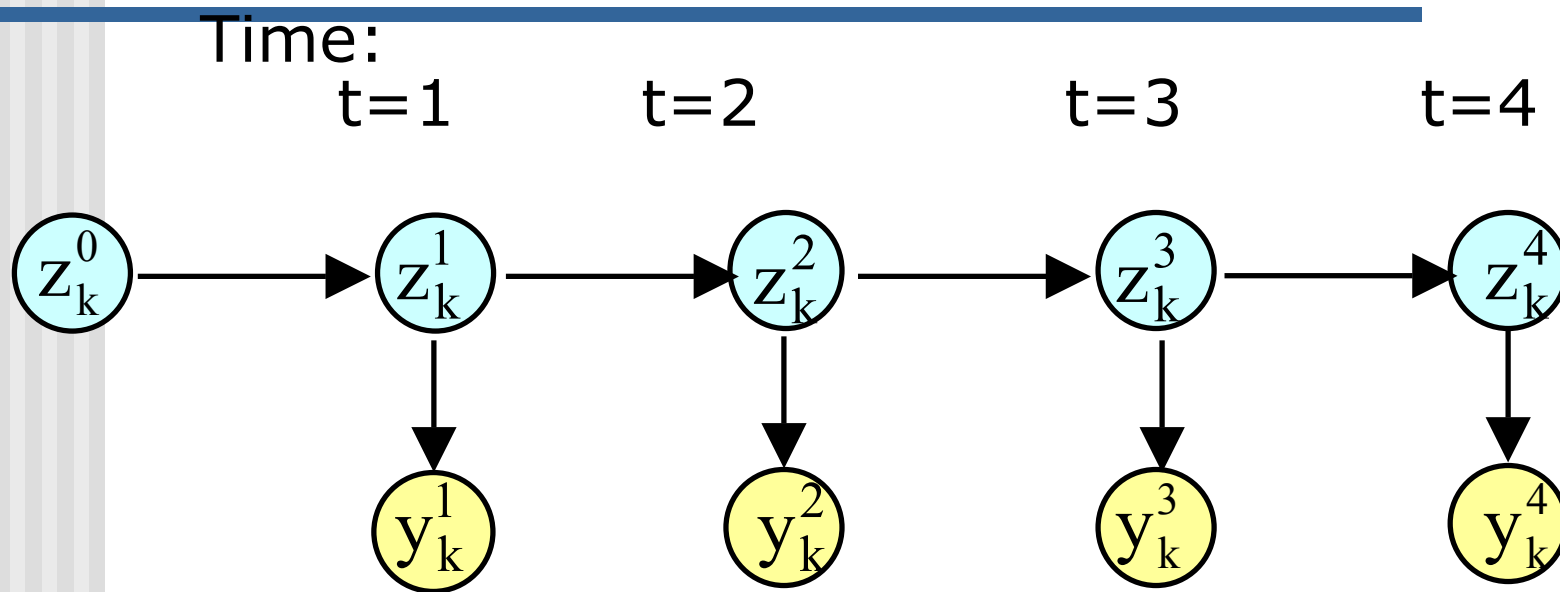
t=4

Transition probability

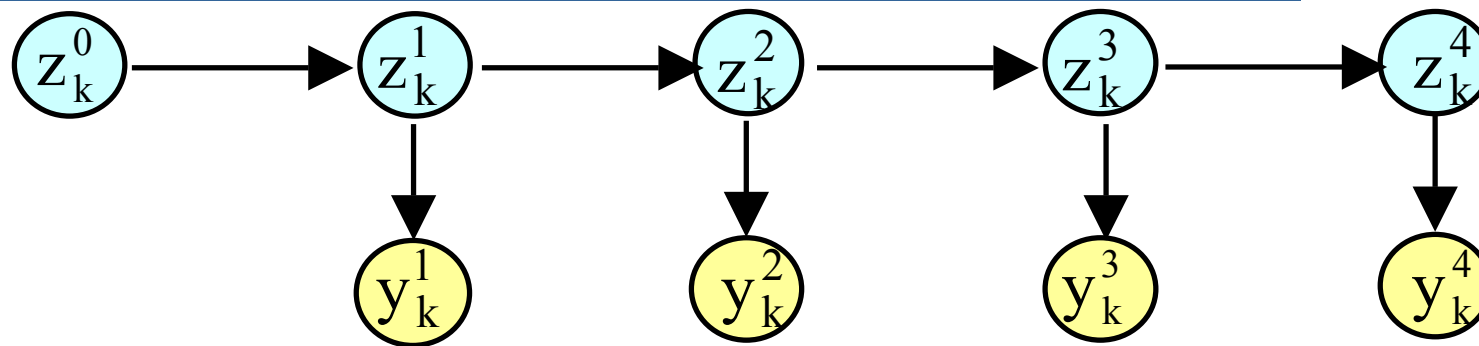
$P(IMI)$



Emission probability



y_k^t = value of biomarker at t^{th} time on k^{th} cow
 $z_k^t = 0$ if IMI- $z_k^t = 1$ if IMI+



$$p(\underline{y}_k) = p(y_k^1, y_k^2, \dots, y_k^T)$$

$$= p(z_k^0 = 0)$$

$$+ \prod_{t=1}^T p(y_k^t | z_k^t = 0) [p(z_k^t = 0 | z_k^{t-1} = 0) + p(z_k^t = 0 | z_k^{t-1} = 1)]$$

$$+ p(y_k^t | z_k^t = 1) [p(z_k^t = 1 | z_k^{t-1} = 0) + p(z_k^t = 1 | z_k^{t-1} = 1)]$$

- Output independence: observations are independent given the unknown IMI state

$$p(y_k^t | z_k^t, y_k^{t-1}, y_k^{t-2}, \dots) = p(y_k^t | z_k^t)$$

- Time is discrete
- Markov property: the next state depends only on the current state

$$p(z_k^{t+1} | z_k^t, z_k^{t-1}, \dots, z_k^1) = p(z_k^{t+1} | z_k^t)$$

“Conditioned on the present, the past & future are independent”

- Stationarity: transition probabilities are time invariant

$$p(z_k^{t+1} = i | z_k^t = j) = a_{ij}$$

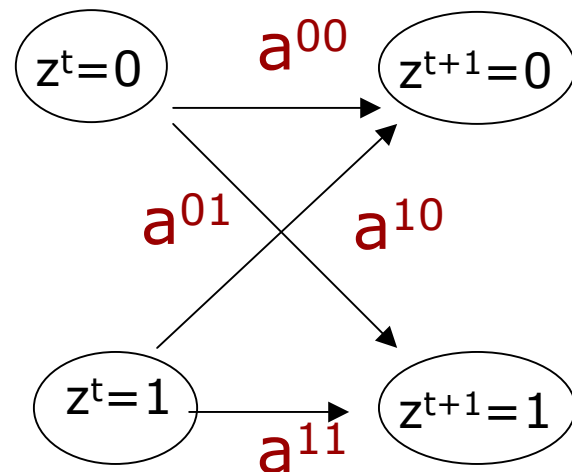
$$H = (A, \mu_0, \mu_1, \sigma_0^2, \sigma_1^2, \lambda_k)$$

$$p(y_k^t | z_k^t = 0) \sim N(\mu_0^t, \sigma_0^2)$$

$$p(z_k^t = 1) \sim \text{Bi}(\lambda_k)$$

$$p(y_k^t | z_k^t = 1) \sim N(\mu_1^t, \sigma_1^2)$$

cows are independent: $p(\underline{y}_1, \underline{y}_2, \dots, \underline{y}_N) = \prod_{k=1}^N p(\underline{y}_k)$



$$p(z_k^t = i | z_k^{t-1} = j) = a_k^{ij}$$

$$\Rightarrow A_k = \begin{pmatrix} a_k^{00} & a_k^{01} \\ a_k^{10} & a_k^{11} \end{pmatrix}$$

3b. Maximum likelihood estimation

Likelihood for one sequence (one cow):

$$p(\underline{y}_k | H) = p(y_k^1, y_k^2, \dots, y_k^T | H) = \sum_{\underline{z}} p(y_k^1, y_k^2, \dots, y_k^T | \underline{z}, H) \times p(\underline{z} | H)$$


$$p(y_k^1, y_k^2, \dots, y_k^T | \underline{z}, H)$$

$$= p(y_k^1, y_k^2, \dots, y_k^T | z_k^1, z_k^2, \dots, z_k^T, H)$$

$$= p(y_k^1 | z_k^1, z_k^2, \dots, z_k^T, H) p(y_k^2 | y_k^1, z_k^1, z_k^2, \dots, z_k^T, H) \dots$$

$$= p(y_k^1 | z_k^1, H) p(y_k^2 | z_k^2, H) \dots p(y_k^T | z_k^T, H)$$

$$= \prod_{t=1}^T p(y_k^t | z_k^t, H)$$

as in FMM

$$p(\underline{z} | H) = p(z_k^0, z_k^1, \dots, z_k^T | H)$$

$$= p(z_k^0 | H) p(z_k^1 | z_k^0, H) p(z_k^2 | z_k^1, z_k^0, H) \dots p(z_k^{T-1} | z_k^{T-2}, \dots, z_k^1, H)$$

$$= p(z_k^0 | H) p(z_k^1 | z_k^0, H) p(z_k^2 | z_k^1, H) \dots p(z_k^{T-1} | z_k^{T-2}, H)$$

$$= p(z_k^0 | H) \prod_{t=1}^T p(z_k^t | z_k^{t-1}, H)$$

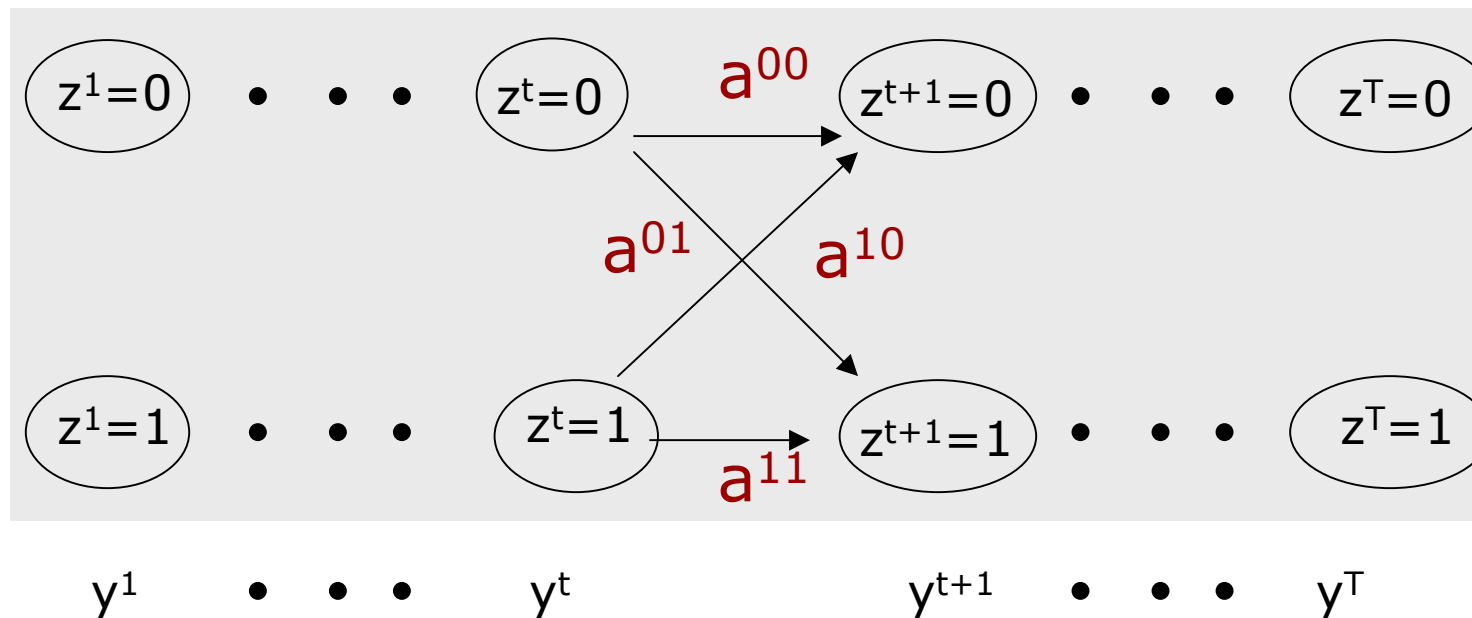
$$p(\underline{y}_k^1, \dots, \underline{y}_k^T | H)$$

$$= \sum_{\underline{z}} p(z_k^0 | H) \prod_{t=1}^T p(z_k^t | z_k^{t-1}, H) p(y_k^t | z_k^t, H)$$

Likelihood for N cows:

$$p(\underline{y} | F) = \prod_{k=1}^N p(\underline{y}_1, \dots, \underline{y}_N | H)$$

Forward-backward algorithm



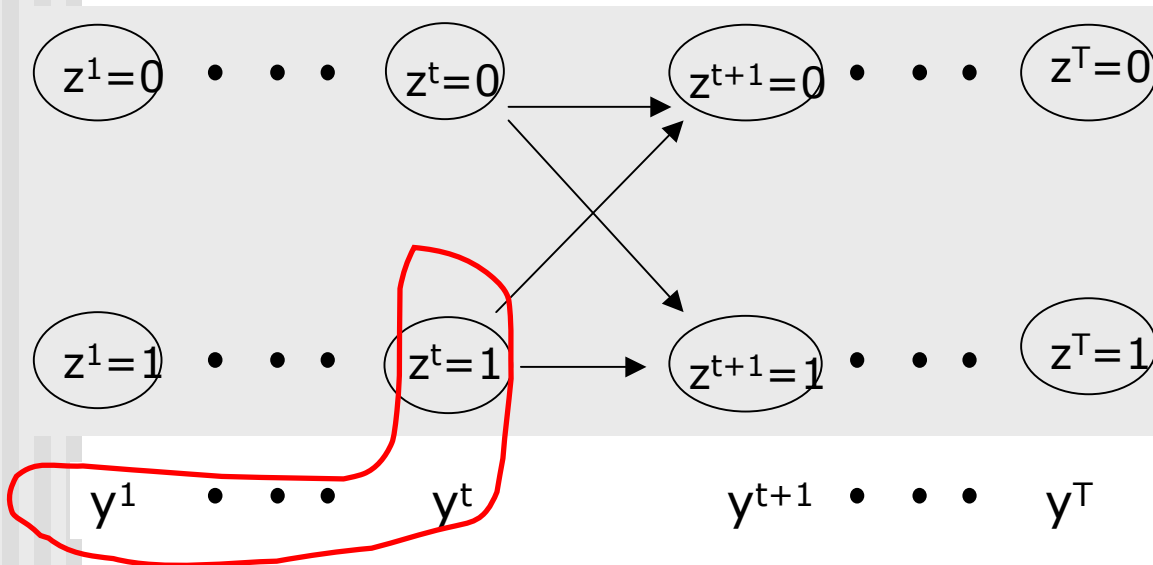
The algorithm takes on the order of $4T$ computations

- **Forward probabilities** : proba that, given H, at time t, the state is i and the sequence of partial observation ($y_1 \dots y_t$) has been generated

$$\alpha_k^t(i) = p(y_k^1 \dots y_k^t, z_k^t = i)$$

$$\alpha_k^0(0) = p(y_k^0 | z_k^0 = 0) p(z_k^0 = 0)$$

$$\alpha_k^{t+1}(0) = [\alpha_k^t(0) a_k^{00} + \alpha_k^t(1) a_k^{01}] p(y_k^t | z_k^t = 0)$$



3 steps

induction ($t=0$)

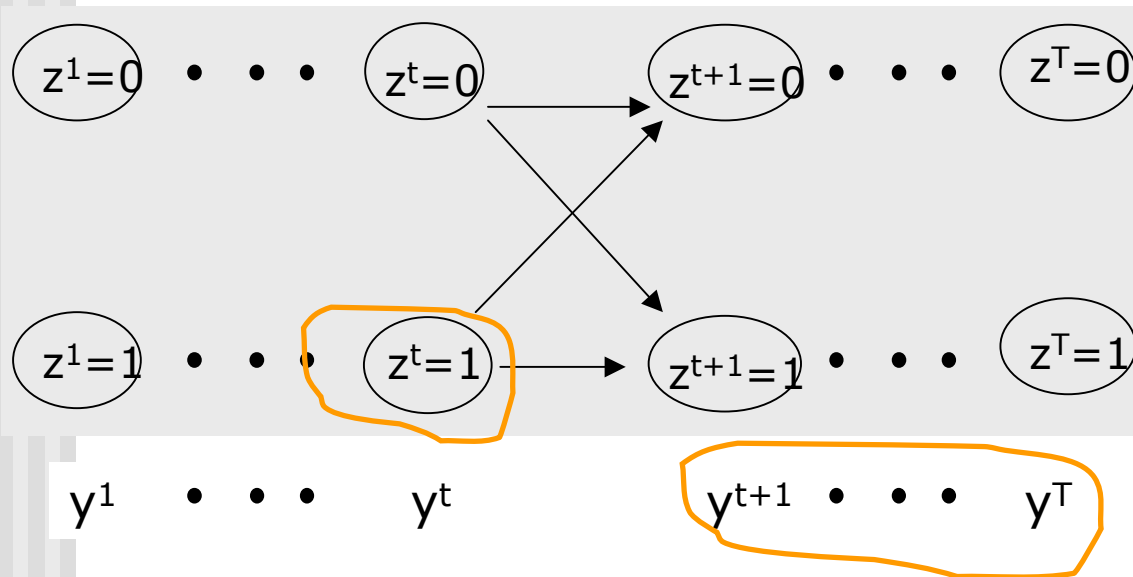
recursion (increasing t)

termination ($t=T$)

- **Backward probabilities** : proba that, given H and given the state i at time t , a sequence of partial observation $(y_{t+1} \dots y_T)$ has been generated

$$\beta_k^t(i) = p(y_k^T \dots y_k^{t+1} \mid z_k^t = i)$$

$$\begin{aligned} \beta_k^t(0) = & [a_k^{00} p(y_k^{t+1} \mid z_k^{t+1} = 0) \beta_k^{t+1}(0)] \\ & + [a_k^{10} p(y_k^{t+1} \mid z_k^{t+1} = 1) \beta_k^{t+1}(1)] \end{aligned}$$

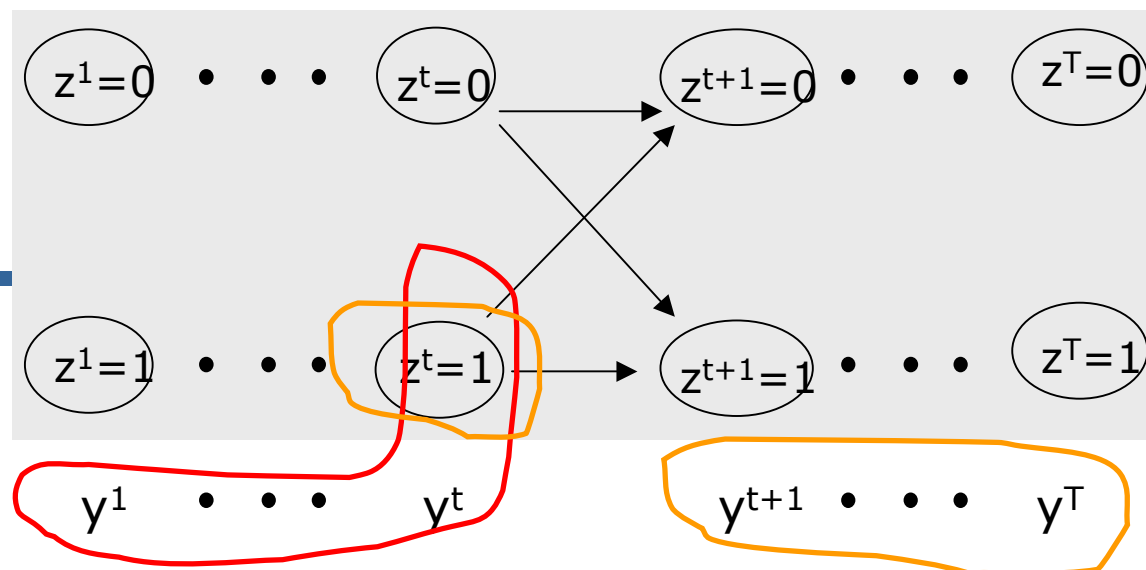


3 steps

induction ($t=T$)

recursion (decreasing t)

termination ($t=0$)



$$\begin{aligned} \alpha_k^t(1) \beta_k^t(1) &= \underline{p(y_k^1 \dots y_k^t, z_k^t = 1 | H)} \quad \underline{p(y_k^{t+1} \dots y_k^T | z_k^t = 1)} \\ &= p(y_k^1 \dots y_k^t, y_k^{t+1} \dots y_k^T, z_k^t = 1 | H) \\ &= p(\underline{y_k}, z_k^t = 1 | H) \end{aligned}$$

$$\alpha_k^t(0) \beta_k^t(0) = p(\underline{y_k}, z_k^t = 0 | H)$$

$$\Rightarrow \alpha_k^t(0) \beta_k^t(0) + \alpha_k^t(1) \beta_k^t(1) = p(\underline{y_k} | H)$$

3c. EM algorithm

E step:

$$Q^{(p-1)} = E[\log[p(\underline{y}_1, \dots, \underline{y}_N; \underline{z}_1, \dots, \underline{z}_N \mid \Theta) \mid \underline{y}_1, \dots, \underline{y}_N; \Theta^{(p-1)}]]$$

$$= \sum_{k=1}^N \left[\begin{aligned} & E[z_k^0 = 1 \mid \underline{y}_k] \log(\lambda_k) + E[z_k^0 = 0 \mid \underline{y}_k] \log(1 - \lambda_k) \\ & + \sum_{t=1}^T \sum_{i,j}^{0,1} E[z_k^t = i, z_k^{t+1} = j \mid \underline{y}_k] \log(a_k^{ij}) \\ & + \sum_{t=1}^T \sum_{i,j}^{0,1} E[z_k^t = i \mid \underline{y}_k] \log N(\mu_i^t, \sigma_i^2) \end{aligned} \right]$$



$$\begin{aligned} E[z_k^t = i \mid \underline{y}_k] &= p[z_k^t = i \mid \underline{y}_k] = \gamma_{i,k}^t \\ &= \frac{p[\underline{y}_k, z_k^t = i]}{p[\underline{y}_k, z_k^t = 0] + p[\underline{y}_k, z_k^t = 1]} \\ &= \frac{\alpha_k^t(i) \beta_k^t(i)}{\alpha_k^t(0) \beta_k^t(0) + \alpha_k^t(1) \beta_k^t(1)} \end{aligned}$$

It is the probability of being in state i at time t , given the observation sequence $y_1 y_2 \dots y_T$.

$$\begin{aligned}
 \rightarrow E[z_k^t = i, z_k^{t+1} = j | \underline{y}_k] &= p[z_k^t = i, z_k^{t+1} = j | \underline{y}_k] = \xi_{ij,k}^t \\
 &= \frac{p[\underline{y}_k, z_k^t = i, z_k^{t+1} = j]}{p[\underline{y}_k, z_k^t = 0] + p[\underline{y}_k, z_k^t = 1]} \\
 &= \frac{\alpha_k^t(i) a_k^{ij} \beta_k^{t+1}(j) N(\mu_j^{t+1}, \sigma_j^2)}{\alpha_k^t(0) \beta_k^t(0) + \alpha_k^t(1) \beta_k^t(1)}
 \end{aligned}$$

It is the probability of being in state i at time t and in state j at time $t+1$, given the observation sequence $y_1 y_2 \dots y_T$.

M step: $\Theta^{(p)} = \operatorname{argmax} Q^{(p-1)}$

$$\hat{\lambda}_k = \gamma_{1,k}^0$$

= expected frequency of state IMI-
at start

$$\hat{a}_{ij}^k = \frac{\sum_{t=1,T} \xi_{ij,k}^t}{\sum_{t=1,T} \gamma_{j,k}^t}$$

= expected number of transitions
from state j to state i divided by
the expected number of transitions
out of state j

= weighted means and
variances

$$\hat{\mu}_i^t = \frac{\sum_{k=1,N} \gamma_{i,k}^t y_k^t}{\sum_{k=1,N} \gamma_{i,k}^t}$$

$$\hat{\sigma}_i^2 = \frac{\sum_{k=1,N} \sum_{t=1,T} \gamma_{i,k}^t (y_k^t - \hat{\mu}_i^t)^2}{\sum_{k=1,N} \sum_{t=1,T} \gamma_{i,k}^t}$$

4. Simulation

Survey of SCC

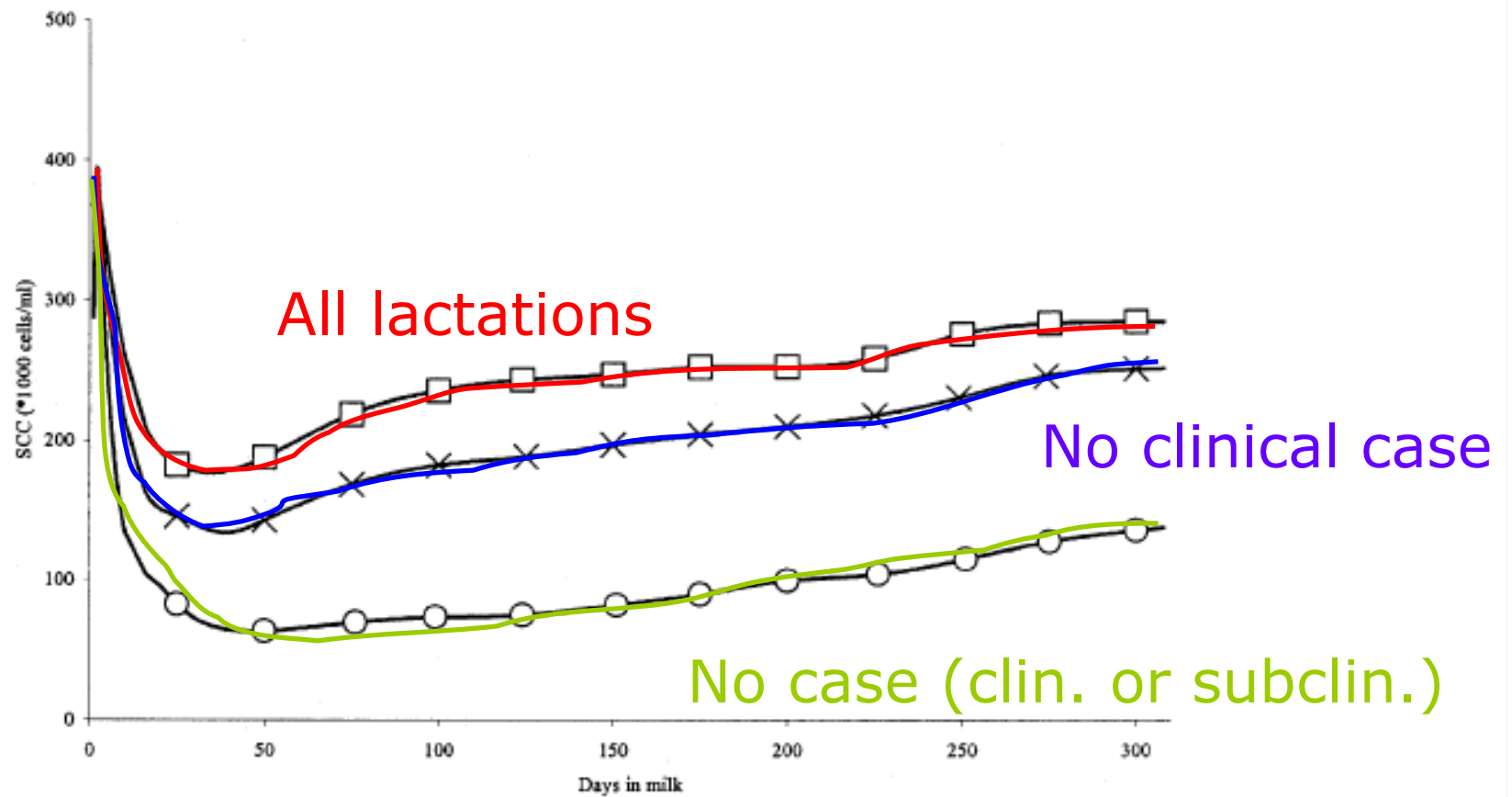
Pathogen

Severity of response

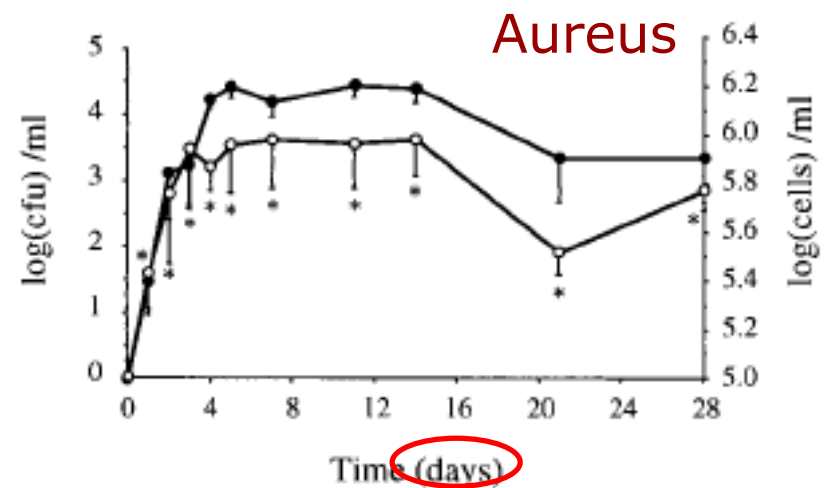
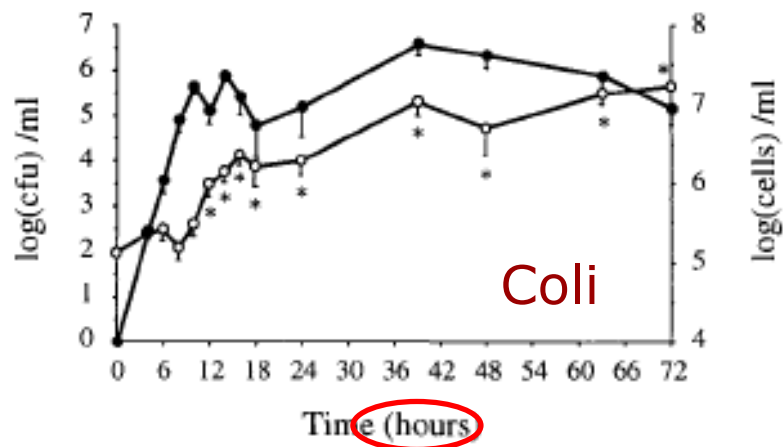
Data sets

Accuracy of MLE

4a. Survey (de Haas et al., 2002, 2004)

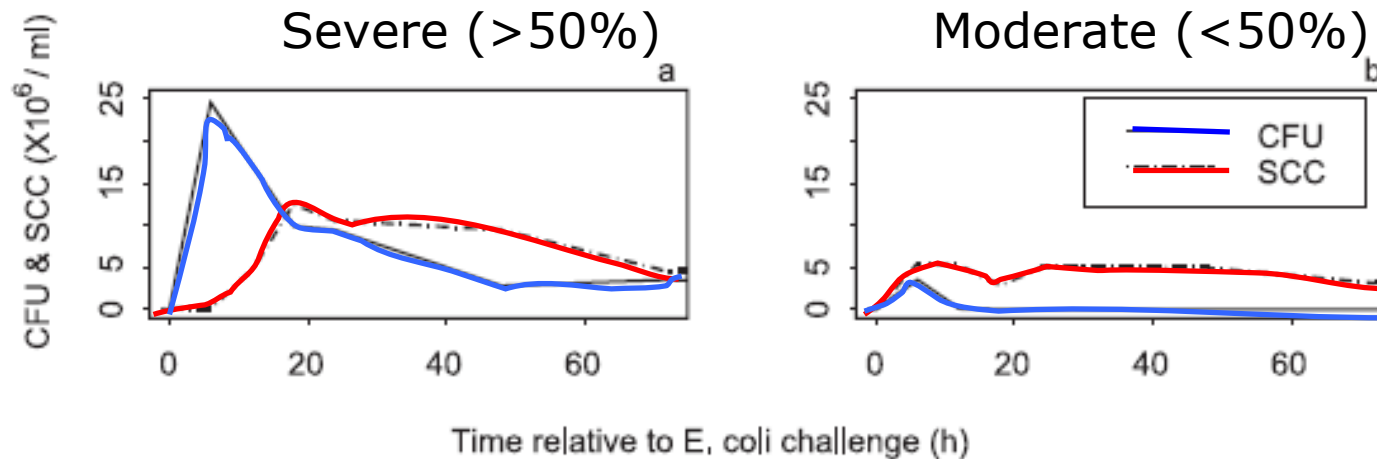


4b. Pathogen



CÉLINE RIOLLET, PASCAL RAINARD,* AND BERNARD POUTREL
CLINICAL AND DIAGNOSTIC LABORATORY IMMUNOLOGY, Mar. 2000, p. 161-167

4c. Severity of response

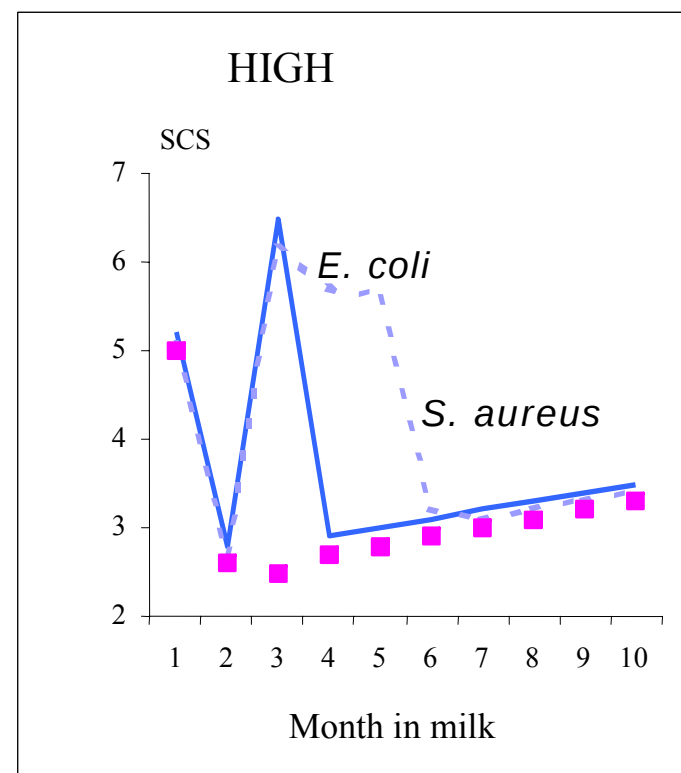
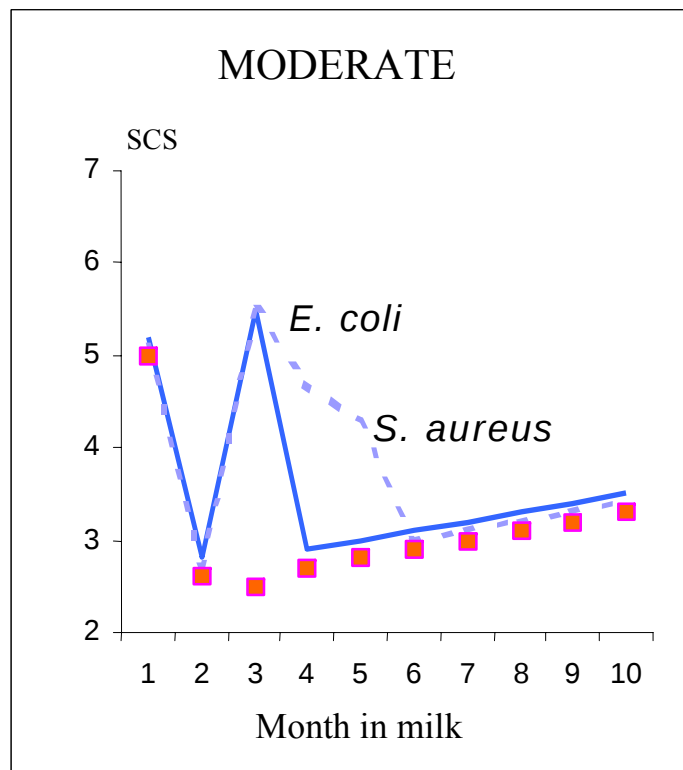


Jalil MEHRZAD^{a,b,c}, Luc DUCHATEAU^a, Christian BURVENICH^{a*}
Vet. Res. 36 (2005) 101–116

4d. Simulated data sets

μ_0^t for $t = 1$ to 10

μ_1^t for $t = 1$ to 10



3 simulated data sets:

% infected cows = 20, 50%

% *E. coli* among infected cows = 0,50,100%

high and moderate responders: μ_1^t

$\sigma_0^2 = 1.0$ or 1.4

EM algorithm for both HMM and FMM:

same priors

500 iterations

10 replications

Accuracy of MLE: $\theta - \hat{\theta}$

4e. Accuracy of MLE

Bias in μ_0^t (3 to 5)

FMM : 0.15 (0.03)

HMM : 0.19 (0.02)

Bias in σ_0^2 (1.0 or 1.4)

FMM : 0.22 (0.04)

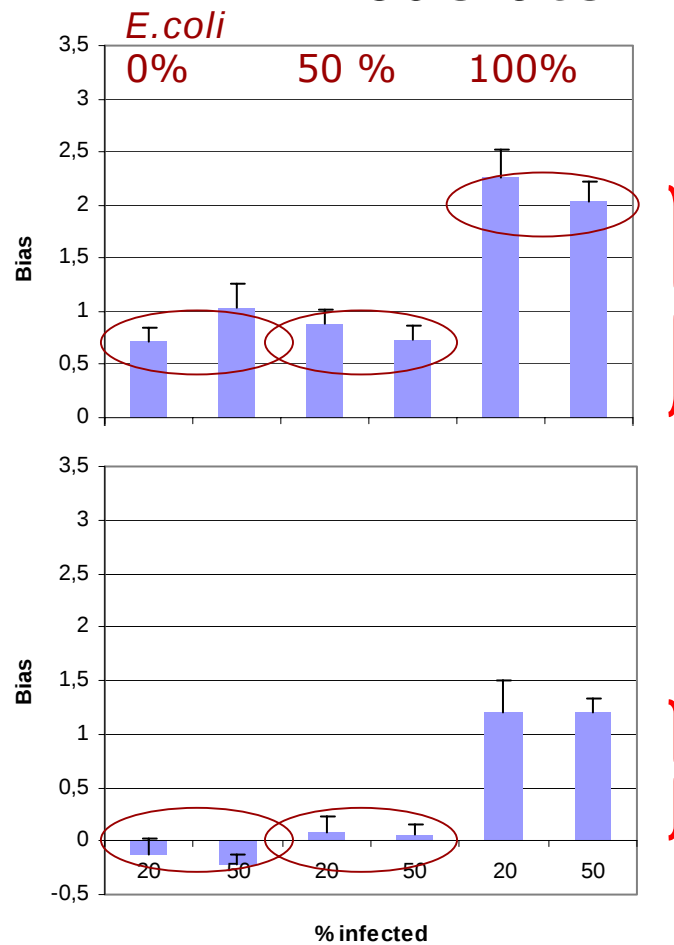
HMM : 0.24 (0.05)

Similar differences between true values and MLE for parameters of the IMI- distribution

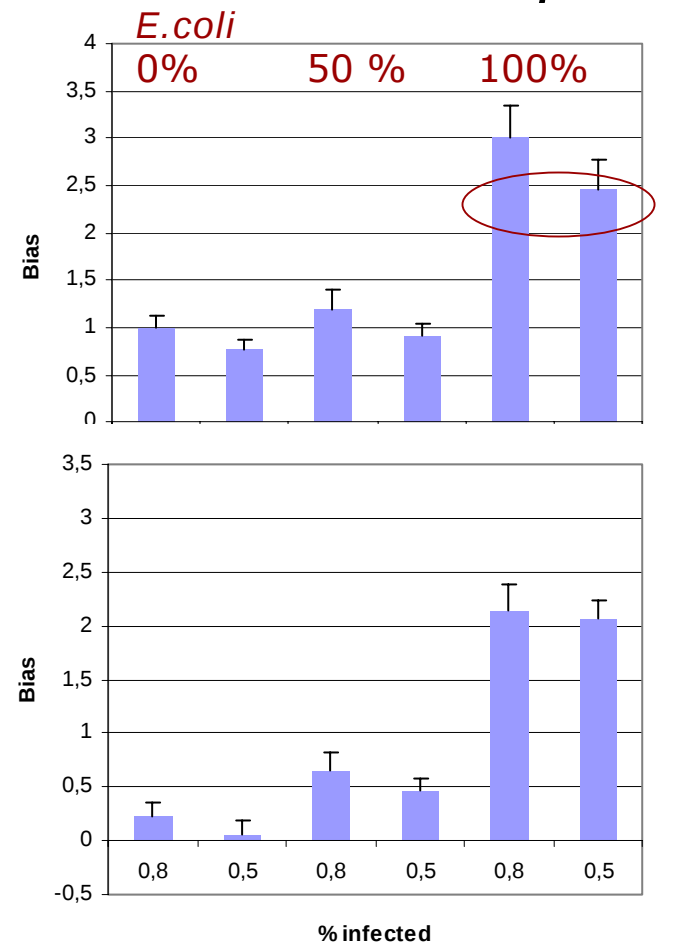
Bias in μ_1^t (5 to 7)

FMM

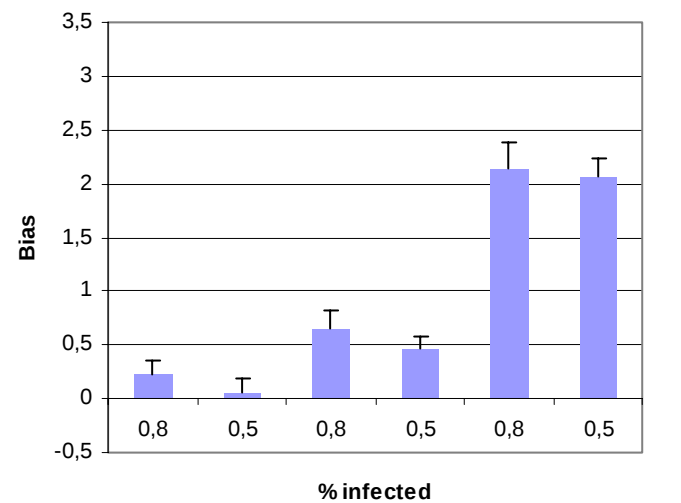
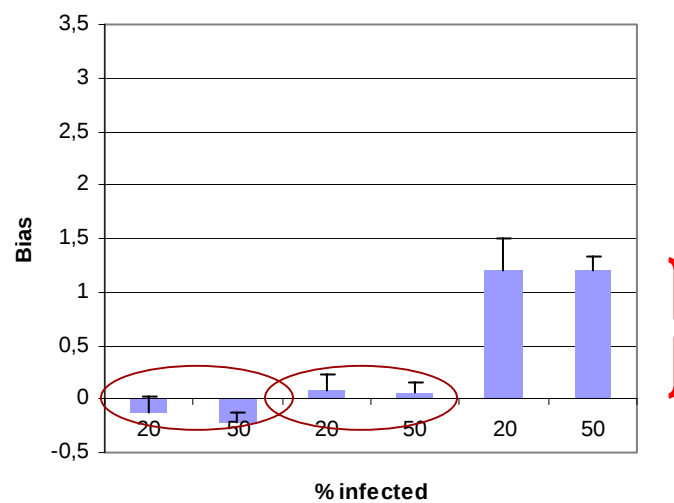
Moderate



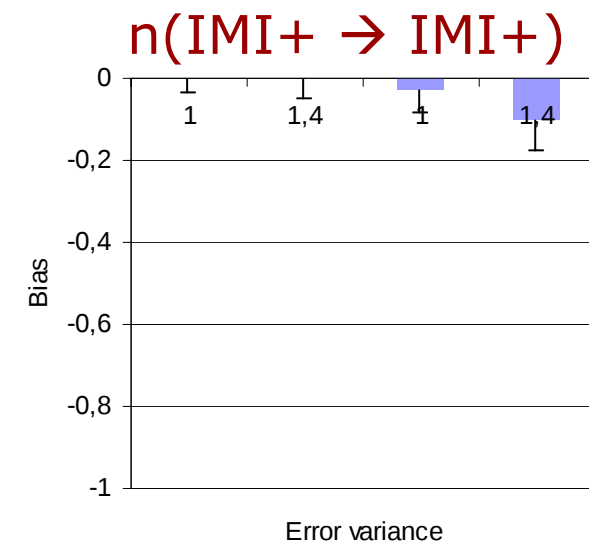
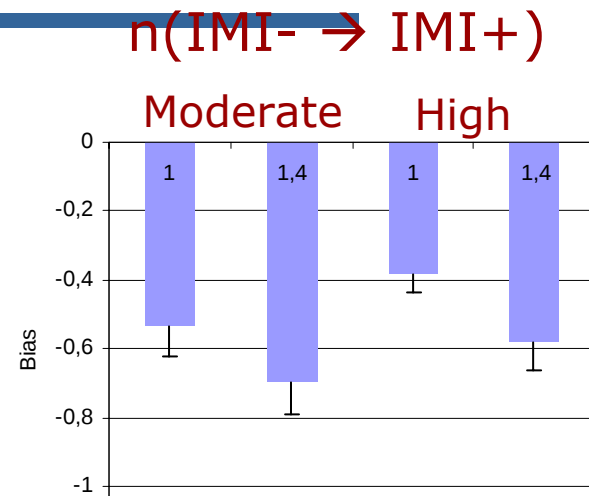
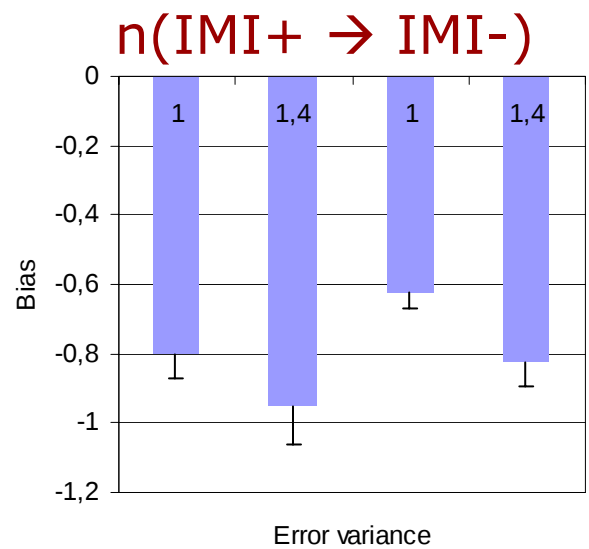
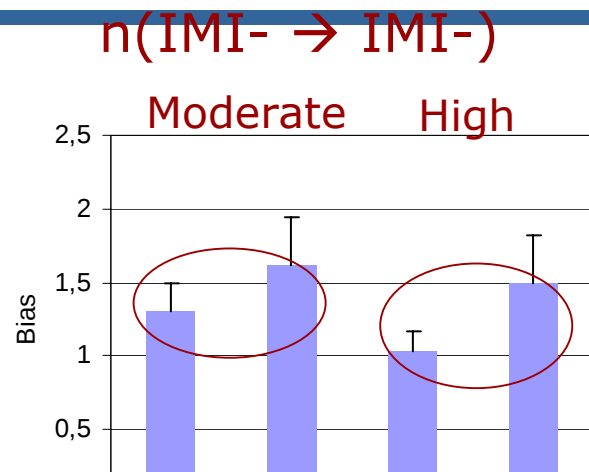
High



HMM



Bias in I



Conclusions

- Same amount of data

- Increased accuracy of MLE
- Resistance and tolerance
- Transition probabilities

- Simplification of reality

- Age, season, herd, ..
- 'Isolated' proba of IMI-
- Genetic relationship between cows





Thank you for your attention